TESTING FOR DENSE SUBSETS IN A GRAPH VIA THE PARTITION FUNCTION*

ALEXANDER BARVINOK † and ANTHONY DELLA PELLA †

Abstract. For a set S of vertices of a graph G, we define its density $0 \le \sigma(S) \le 1$ as the ratio of the number of edges of G spanned by the vertices of S to $\binom{|S|}{2}$. We show that, given a graph G with n vertices and an integer $m \ll n$, the partition function $\sum_{S} \exp\{\gamma m \sigma(S)\}$, where the sum is taken over all m-subsets S of vertices and $0 < \gamma < 1$ is fixed in advance, can be approximated within relative error $0 < \epsilon < 1$ in quasi-polynomial $n^{O(\ln m - \ln \epsilon)}$ time. We discuss numerical experiments and observe that for the random graph G(n, 1/2) one can afford a much larger γ , provided the ratio n/m is sufficiently large.

Key words. graph, density, partition function, algorithm, complex zeros

AMS subject classifications. 05C31, 82B20, 05C85, 05C69, 68Q25

DOI. 10.1137/19M1247413

1. Introduction and main results. Let G = (V, E) be an undirected graph, without loops or multiple edges. For a nonempty subset $S \subset V$ of vertices, we define the *density* $\sigma(S)$ as the fraction of the pairs of vertices of S that span an edge of G:

$$\sigma(S) = \frac{\left|\binom{S}{2} \cap E\right|}{\binom{|S|}{2}},$$

where $\binom{S}{2}$ is the set of all unordered pairs of vertices from S. Hence $0 \le \sigma(S) \le 1$ for all subsets, $\sigma(S) = 0$ if S is an *independent set* and $\sigma(S) = 1$ if S is a *clique*.

We are interested in the following general problem: given a graph G = (V, E)with |V| = n vertices and an integer $m \leq n$, estimate the highest density of an *m*subset $S \subset V$. This is, of course, a hard problem: for example, testing whether a given graph contains a clique of a given size, or even estimating the size of the largest clique within a factor of $n^{1-\epsilon}$ for any $\epsilon > 0$, fixed in advance, is already an NPhard problem [12], [15]. Moreover, modulo some plausible complexity assumptions, it is hard to approximate the highest density of an *m*-subset for a given *m*, within a constant factor, fixed in advance [5]. The best known efficient approximation achieves the factor of $n^{1/4}$ in quasi-polynomial $n^{O(\ln n)}$ time [6]. There are indications that the factor $n^{1/4}$ might be hard to beat [7]. We note that the most interesting case is when *m* grows and $n \gg m$, since the highest density of an *m*-subset can be computed in polynomial time up to an additive error of $\epsilon n^2/m^2$ for any $\epsilon > 0$, fixed in advance [10] (and if *m* is fixed in advance, the densest *m*-subset can be found by the exhaustive search in polynomial time).

1.1. Partition function. In this paper, we approach the problem of finding the densest, or just a reasonably dense subset, via computing the *partition function*

^{*}Received by the editors February 28, 2019; accepted for publication (in revised form) November 27, 2019; published electronically January 30, 2020.

https://doi.org/10.1137/19M1247413

Funding: The first author was partially supported by NSF grants DMS 1361541 and DMS 1855428.

[†]Department of Mathematics, University of Michigan, Ann Arbor, MI 48109-1043 (barvinok@ umich.edu, adellape@umich.edu).

(1.1)
$$\operatorname{den}_{m}(G;\gamma) = \binom{n}{m}^{-1} \sum_{\substack{S \subset V:\\|S|=m}} \exp\left\{\gamma m \sigma(S)\right\},$$

where $\gamma > 0$ is a parameter. We are interested in computing (approximating) den_m(G; γ) efficiently. The *exponential tilting*, $\sigma(S) \mapsto \exp{\{\gamma m \sigma(S)\}}$ (see, for example, section 13.7 of [14]), puts greater emphasis on the sets of higher density. Let us consider the set $\binom{V}{m}$ of all *m*-subsets of V as a probability space with the uniform measure. By the Markov inequality, for any $0 < \sigma_0 < 1$, we have

(1.2)
$$\sigma_0 + \frac{\ln \mathbf{P}(\sigma(S) \ge \sigma_0)}{\gamma m} \le \frac{\ln \operatorname{den}_m(G;\gamma)}{\gamma m} \le \max_{\substack{S \subset V:\\|S| = m}} \sigma(S),$$

so the larger γ we can afford, the better approximation for the densest *m*-subset we get. In particular, if we could choose $\gamma \gg \ln n$, then from (1.2) we could approximate the highest density of an *m*-subset within an arbitrarily small additive error.

The partition function (1.1) was introduced in [3], where an algorithm of quasipolynomial $n^{O(\ln m - \ln \epsilon)}$ complexity was constructed to compute (1.1) within relative error $0 < \epsilon < 1$, when $\gamma = 0.07$ and when $\gamma = 0.27$, under additional assumptions that $n \ge 8m$ and $m \ge 10$. It follows from (1.2) that if the probability to hit an *m*-subset *S* of density at least σ_0 at random is $e^{-o(m)}$, then we can certify the existence of an *m*-subset of density at least $\sigma_0 - o(1)$ in quasi-polynomial time, just by computing (1.1). It is also shown in [3] that by successive conditioning, one can find in quasipolynomial time an *m*-subset *S* with density at least as high as certified by the value of (1.1).

In this paper, we present an algorithm, which, for any $0 < \gamma < 1$, fixed in advance, and a given $0 < \epsilon < 1$, computes the value of (1.1) within relative error ϵ in quasi-polynomial $n^{O(\ln m - \ln \epsilon)}$ time, provided $n > \omega(\gamma)m$ for some constant $\omega(\gamma) > 1$. This improvement from $\gamma = 0.27$ to $\gamma = 1$ makes the algorithm competitive in some situations where it was not competitive before. Suppose, for example, we want to separate efficiently the graphs that have sufficiently many *m*-cliques from the graphs that are sufficiently far from having a single *m*-clique. Below we show that for $\gamma < 0.5$ our algorithm is inferior to a simple test based on the Kruskal–Katona Theorem, while for $\gamma > 0.5$ the former can cover a greater range than the latter.

Example 1 (Testing graphs for *m*-cliques). Let us fix two numbers $0 < \delta < 1$ and $\alpha > 0$ and consider the following two mutually exclusive conditions.

CONDITION 1.1. For every $S \subset V$ such that |S| = m we have $\sigma(S) \leq 1 - \delta$ and

CONDITION 1.2. If $S \subset V$ is a random subset, sampled uniformly from the set $\binom{V}{m}$ of all m-sets of vertices, then the probability that S is a clique is at least $e^{-\alpha m}$.

Suppose further, we are presented with a graph G = (V, E) and told that either Condition 1.1 or Condition 1.2 holds. Our goal is to decide which one. This is somewhat in the spirit of "property testing" [11].

We observe that if Condition 1.1 holds, then $\operatorname{den}_m(G;\gamma) \leq e^{\gamma m(1-\delta)}$ and if Condition 1.2 holds, then $\operatorname{den}_m(G;\gamma) \geq e^{(\gamma-\alpha)m}$. Consequently, if

$$(1.3) \qquad \qquad \alpha < \gamma \delta$$

and we can approximate $den_m(G; \gamma)$ efficiently, we can efficiently tell Condition 1.1 and Condition 1.2 apart. An anonymous referee to [3] noticed that another, much simpler, algorithm can be inferred from the Kruskal–Katona Theorem. Let |V| = n. If Condition 1.1 holds, then $|E| \leq (1-\delta)\binom{n}{2}$. The Kruskal–Katona Theorem (see, for example, section 5 of [8]) implies that if Condition 1.2 holds, then for every k such that $\binom{k}{m} \leq e^{-\alpha m} \binom{n}{m}$, we must have $|E| \geq \binom{k}{2}$, the model case being a graph G consisting of a k-clique and n-k isolated vertices. A computation shows that as $n \to \infty$, we can tell Condition 1.1 and Condition 1.2 apart just by counting the edges of G, provided

(1.4)
$$\alpha < -\frac{1}{2}\ln(1-\delta).$$

Comparing (1.3) and (1.4), we observe that the algorithm based on computing the partition function den_m(G; γ) is not competitive as long as $\gamma < 0.5$, which is the case in [3], but becomes competitive at least for small values of δ as soon as $\gamma > 0.5$. Numerical estimates show that as long as we can choose $\gamma > 0$ arbitrarily close 1, the condition (1.3) serves a wider range of α than the condition (1.4) provided $\delta < 0.7968$.

We still don't know, however, if (1.1) can be efficiently computed for any $\gamma > 0$, fixed in advance, and as we remarked above, it is unlikely that (1.1) can be efficiently computed for $\gamma \gg \ln n$. Our numerical experiments seem to indicate that we can afford a substantially larger γ . This can be partially explained by the fact that for the Erdős–Rényi random graph G(n, 0.5) indeed a much larger γ can be used with high probability; see Theorem 1.4 below.

The improvement from $\gamma = 0.27$ to an arbitrary $\gamma < 1$ required the addition of some new ideas to the technique of [3]. The approach of [3] and of this paper are based on the "interpolation method" [4]. As applied to our case, the idea of the method is to consider den_m(G; z) for a *complex* parameter z. We can efficiently approximate den_m(G; z) at $z = \gamma$ if there is a connected open set $U \subset \mathbb{C}$, not dependent on m or G, such that $0 \in U$, $\gamma \in U$ and den_m(G; z) $\neq 0$ for all $z \in U$. In [3], the set U is a disc centered at z = 0, whereas in the current paper it is a thin neighborhood of the interval $[0, \gamma]$, which allows us to reach larger γ , but also requires a more refined analysis to establish zero-freeness. We give some more details now.

1.2. Multivariate partition function. Given $n \times n$ symmetric complex matrix $Z = (z_{ij})$ and $2 \le m \le n$, we define

(1.5)
$$P_m(Z) = \sum_{\substack{S \subset \{1,...,n\} \\ |S|=m}} \exp\left\{\sum_{\substack{\{i,j\} \subset S \\ i \neq j}} z_{ij}\right\}.$$

Note that the diagonal entries of Z are irrelevant, so we assume that $z_{ii} = 0$ for all *i*. Given a graph G = (V, E) with set $V = \{1, \ldots, n\}$ of vertices and $\gamma > 0$, we

define $Z_0 = (z_{ij})$ by $(z_{ij}) = (z_{ij}) = (z_{ij})$

$$z_{ij} = \begin{cases} \frac{\gamma}{m-1} & \text{if } \{i,j\} \in E\\ -\frac{\gamma}{m-1} & \text{if } \{i,j\} \notin E \end{cases}$$

and observe that

(1.6)
$$P_m(Z_0) = \sum_{\substack{S \subset \{1, \dots, n\} \\ |S| = m}} \exp\left\{m\gamma\sigma(S) - \frac{\gamma m}{2}\right\}$$
$$= \exp\left\{-\frac{\gamma m}{2}\right\} \binom{n}{m} \operatorname{den}_m(G;\gamma).$$

Hence to compute (1.1) it suffices to compute $P_m(Z_0)$. We compute $P_m(Z_0)$ by interpolation; see [3], [4]. For that, it suffices to show that $P_m(Z) \neq 0$ in some neighborhood of a path connecting the zero matrix to Z_0 in the space of complex matrices.

We prove the following result.

THEOREM 1.3. For any $0 < \delta < 1$ there exist $\eta = \eta(\delta) > 0$ and $\omega = \omega(\delta) > 1$ such that if $n \ge \omega m$, then $P_m(Z) \ne 0$ for any $n \times n$ symmetric complex matrix $Z = (z_{ij})$ such that

$$|\Re z_{ij}| \leq \frac{\delta}{m-1}$$
 and $|\Im z_{ij}| \leq \frac{\eta}{m-1}$ for all $1 \leq i \neq j \leq n$.

We prove Theorem 1.3 in sections 2 and 3. Using Theorem 1.3, in section 4 we present an algorithm of quasi-polynomial $n^{O(\ln m)}$ complexity to compute $P_m(Z_0)$ and hence den_m(G; γ) for any $0 < \gamma < 1$, fixed in advance.

In [3] it was established that $P_m(Z) \neq 0$ in a polydisc

$$\mathcal{D}_{m,n} = \left\{ Z = (z_{ij}) : |z_{ij}| \le \frac{0.27}{m-1} \quad \text{for all} \quad 1 \le i \ne j \le n \right\}$$

provided $n \gg m$ and m is large enough. In Theorem 1.3, we establish that $P_m(Z) \neq 0$ in a more "economical" domain, "stretched" along the real part of the complex space of matrices. This allows us to improve the constant γ for which den_m(G; γ) is still efficiently computable.

In section 5, we discuss some results of our numerical experiments, which seem to indicate that we can afford an essentially bigger δ in Theorem 1.3. This can be partially explained by the fact that for the Erdős–Rényi random graph G(n, 0.5) this is indeed the case. Namely, we prove the following result in section 6.

THEOREM 1.4. Let us choose positive integers n and $2 \le m \le n$. For $n \times n$ symmetric matrix $W = (w_{ij})$ of independent random variables, where

$$\mathbf{P}(w_{ij} = 1) = \mathbf{P}(w_{ij} = -1) = \frac{1}{2},$$

we define the polynomial

$$h_W(z) = \binom{n}{m}^{-1} \sum_{\substack{S \subset \{1, \dots, n\} \\ |S| = m}} \prod_{\{i, j\} \subset S} (1 + zw_{ij}).$$

Let r > 0 and $\tau > 1$ be real numbers. If $n \ge 2m^2 (1+r^2)^m + 2m$, then the probability that $h_W(z)$ has a root in the disc $|z| < r/\sqrt{2\tau}$ does not exceed $1/\tau$.

In particular, if $n \gg m^2$, then with high probability $h_W(z)$ has no roots in the disc $|z| < c/\sqrt{m}$, for an arbitrary large c > 0, fixed in advance. Similarly, if $\ln n \gg m$, then with high probability $h_W(z)$ has no roots in the disc |z| < c for an arbitrary large c > 0, fixed in advance.

The polynomial $h_W(z)$ is easily translated into the partition function $den_m(G; \gamma)$, where G is the graph with set $V = \{1, \ldots, n\}$ of vertices and two vertices $\{i, j\}$ span an edge if and only if $w_{ij} = 1$: for $0 < \alpha < 1$, we have

(1.7)
$$h_W(\alpha) = (1-\alpha)^{\binom{m}{2}} \operatorname{den}_m(G;\gamma) \quad \text{where} \quad \gamma = \frac{m-1}{2} \ln \frac{1+\alpha}{1-\alpha}.$$

Consequently, with high probability we can can approximate $\operatorname{den}_m(G;\gamma)$ in quasipolynomial time for γ as large as $\gamma = \sqrt{m}$ provided $n \gg m^2$ and as large as $\gamma = m$ provided $\ln n \gg m$. Since the graphs we experimented on were to a large degree random (but not necessarily Erdős–Rényi G(n, 0.5)), we may have obtained overly optimistic numerical evidence.

As is easily seen, $\mathbf{E} h_W(\alpha) = 1$ and from our proof in section 6 it follows that $h_W(\alpha)$ is strongly concentrated. For example, in the regime of $n = \Omega(m^2)$ and $\alpha = 1/\sqrt{m}$, we have **var** $h_W(\alpha) = O(1)$. This concentration, however, does not allow us to predict with high probability the value of $h_W(\alpha)$ with the precision that the interpolation technique based on Theorem 1.4 allows for.

In section 6, we also discuss what may happen if G is a random graph G(n, 0.5) with a planted *m*-clique.

2. Preliminaries. We consider the partition function P_m of section 1.2 within a family of partition functions, which will allow us to prove Theorem 1.3 by induction.

2.1. Functionals $P_{\Omega}(Z)$. Let us fix integers n and $2 \le m \le n$. For a subset $\Omega \subset \{1, \ldots, n\}$ and $n \times n$ complex symmetric matrix $Z = (z_{ij})$, we define

$$P_{\Omega}(Z) = \sum_{\substack{S \subset \{1, \dots, n\}:\\|S|=m, \Omega \subset S}} \exp \left\{ \sum_{\substack{\{i, j\} \subset S\\i \neq j}} z_{ij} \right\},$$

where we agree that $P_{\Omega}(Z) = 0$ if $|\Omega| > m$. In other words, we restrict the sum (1.5) defining $P_m(Z)$ onto subsets S containing a given set Ω . In particular,

$$P_{\Omega}(Z) = P_m(Z)$$
 if $\Omega = \emptyset$.

The induction will be built on the following straightforward formulas:

(2.1)
$$P_{\Omega}(Z) = \frac{1}{m - |\Omega|} \sum_{j \in \{1, \dots, n\} \setminus \Omega} P_{\Omega \cup \{j\}}(Z) \text{ provided } |\Omega| < m$$

and for $i \neq j$, we have

(2.2)
$$\frac{\partial}{\partial z_{ij}} P_{\Omega}(Z) = \begin{cases} P_{\Omega}(Z) & \text{if } i, j \in \Omega, \\ P_{\Omega \cup \{j\}}(Z) & \text{if } i \in \Omega, j \notin \Omega, \\ P_{\Omega \cup \{i\}}(Z) & \text{if } i \notin \Omega, j \in \Omega, \\ P_{\Omega \cup \{i,j\}}(Z) & \text{if } i, j \notin \Omega. \end{cases}$$

We will often consider complex numbers as vectors in the plane, by identifying $\mathbb{C} = \mathbb{R}^2$ and measuring, in particular, angles between nonzero complex numbers. We will use the following geometric lemma.

LEMMA 2.1. Let $u_1, \ldots, u_n \in \mathbb{C}$ be nonzero complex numbers such that the angle between any two does not exceed θ for some $0 < \theta < \pi/2$. Suppose that

$$\Im\left(\sum_{j=1}^{n} u_j\right) = 0 \quad and \quad \sum_{j=1}^{n} |u_j| = c.$$

Then

$$\sum_{j=1}^{n} |\Im u_j| \leq c \sin \frac{\theta}{2}.$$

Proof. Scaling u_j , if necessary, without loss of generality we assume that c = 1. Without loss of generality, we assume that $\arg u_j \neq 0$ for j = 1, ..., n. Indeed, if $\arg u_j = 0$ for some j, we can remove the vector from the collection, which would make the sum

$$(2.3) \qquad \qquad \sum_{j=1}^{n} |u_j|$$

only smaller. Rescaling $u_j \mapsto \tau u_j$ for some real $\tau > 1$, we make (2.3) equal to 1 and increase

(2.4)
$$\sum_{j=1}^{n} |\Im u_j|$$

Reflecting the vectors u_j in the coordinate axes if necessary, without loss of generality we may assume that $\Re u_1 \ge 0$ and $\Im u_1 > 0$. Hence there is a vector, say u_2 , such that $\Im u_2 < 0$. We necessarily have $\Re u_2 \ge 0$, since otherwise the angle between u_1 and u_2 exceeds $\pi/2$. Then for any vector u_j , we must have $\Re u_j \ge 0$, since otherwise one of the angles formed by u_j with u_1 or u_2 will exceed $\pi/2$.

Hence without loss of generality, we assume that $\Re u_j > 0$ for $j = 1, \ldots, n$. Let

$$\alpha = \max_{j=1,\dots,n} \arg u_j,$$

so that

 $0 \ < \ \alpha \ < \ \theta,$

and let

$$-\beta = \min_{j=1,\dots,n} \arg u_j < 0.$$

Then $\alpha + \beta \leq \theta$.

Let

$$J_{+} = \{j : \arg u_j > 0\}$$
 and $J_{-} = \{j : \arg u_j < 0\}$

Next, without loss of generality, we assume that $\arg u_j = \alpha$ for all $j \in J_+$ and that $\arg u_j = -\beta$ for all $j \in J_-$. Indeed, suppose that $\arg u_1 = \alpha_1$, where $0 < \alpha_1 < \alpha$. We can modify

$$u_1 \longmapsto \frac{\sin \alpha_1}{\sin \alpha} e^{i(\alpha - \alpha_1)} u_1$$

(we rotate and shrink u_1 so as to make its argument equal to α and leave $\Im u_1$ intact). The sum (2.3) gets smaller while all other conditions and the sum (2.4) remain intact. Rescaling $u_j \mapsto \tau u_j$ for some real $\tau > 1$, we make (2.3) equal to 1 and increase (2.4), while keeping other constraints of the lemma intact. The case of arg $u_j > -\beta$ for some $j \in J_-$ is handled similarly.

Next, without loss of generality, we assume that $\alpha + \beta = \theta$. Indeed, if $\alpha + \beta < \theta$, we can rotate and scale vectors u_j as above, so that the sum (2.4) increases while all other conditions are satisified.

Now, let

$$u_+ = \sum_{j \in J_+} u_j$$
 and $u_- = \sum_{j \in J_-} u_j$

Then $\arg u_+ = \alpha$, $\arg u_- = -\beta$, $\Im (u_+ + u_-) = 0$, $|u_+| + |u_-| = 1$, and (2.4) is equal to $|\Im u_+| + |\Im u_-|$.

Denoting $a = |u_+|$ and $b = |u_-|$, we have a + b = 1 and $a \sin \alpha - b \sin \beta = 0$, from which

$$a = \frac{\sin \beta}{\sin \alpha + \sin \beta}$$
 and $b = \frac{\sin \alpha}{\sin \alpha + \sin \beta}$,

and so

$$|\Im u_+| + |\Im u_-| = \frac{2\sin\alpha\sin\beta}{\sin\alpha + \sin\beta}.$$

Now, the function

$$\alpha \longmapsto \frac{1}{\sin \alpha} \quad \text{for} \quad 0 \le \alpha \le \frac{\pi}{2}$$

is convex and hence the minimum of

$$\frac{\sin\alpha+\sin\beta}{\sin\alpha\sin\beta}=\frac{1}{\sin\alpha}+\frac{1}{\sin\beta}$$

on the interval $\alpha + \beta = \theta$, $\alpha, \beta \ge 0$, is attained at $\alpha = \beta = \theta/2$. The proof now follows.

We need another geometric lemma.

LEMMA 2.2. Let $u_1, \ldots, u_n \in \mathbb{C}$ be nonzero complex numbers such that the angle between any two does not exceed θ for some $0 \leq \theta < 2\pi/3$. Let $u = u_1 + \cdots + u_n$. Then

$$|u| \geq \left(\cos\frac{\theta}{2}\right) \sum_{k=1}^{n} |u_k|.$$

Proof. This is Lemma 3.1 of [3] and Lemma 3.6.3 of [4].

3. Proof of Theorem 1.3. We identify the space of $n \times n$ zero-diagonal complex symmetric matrices $Z = (z_{ij})$ with $\mathbb{C}^{\binom{n}{2}}$. Given $\delta \geq \eta > 0$, we define a domain $\mathcal{U}(\delta, \eta) = \mathcal{U}_{n,m}(\delta, \eta) \subset \mathbb{C}^{\binom{n}{2}}$ by

$$\mathcal{U}(\delta,\eta) = \left\{ Z = (z_{ij}) : |\Re z_{ij}| \leq \frac{\delta}{m-1} \text{ and } |\Im z_{ij}| \leq \frac{\eta}{m-1} \right\}.$$

If $Z' = (z'_{ij})$ and $Z'' = (z''_{ij})$ are two matrices from $\mathcal{U}(\delta, \tau)$, then

$$|z'_{ij} - z''_{ij}| \leq \frac{\sqrt{(2\delta)^2 + (2\eta)^2}}{m-1} \leq \frac{2\sqrt{2\delta}}{m-1}$$
 for all i, j .

We will prove by descending induction on $|\Omega|$ that $P_{\Omega}(Z) \neq 0$ for all $Z \in \mathcal{U}(\delta, \eta)$ and that, moreover, a number of stronger conditions are met. The induction is based on the following two lemmas that describe how $P_{\Omega}(Z)$ changes when only the entries in the *i*th row and column of Z change. The first lemma deals with the case of $i \in \Omega$.

LEMMA 3.1. Let us fix $\Omega \subset \{1, \ldots, n\}$ such that $|\Omega| < m$. Suppose that for any $Z \in \mathcal{U}(\delta, \eta)$ and any $j, k \notin \Omega$, we have $P_{\Omega \cup \{j\}}(Z) \neq 0$, $P_{\Omega \cup \{k\}}(Z) \neq 0$ and the angle between the two nonzero complex numbers does not exceed θ for some $0 < \theta \leq \pi/2$. Then

• Part 1: We have

$$P_{\Omega}(Z) \neq 0$$
 for all $Z \in \mathcal{U}(\delta, \eta)$.

314

• Part 2: Suppose additionally that $\Omega \neq \emptyset$, and let us fix an $i \in \Omega$. Let $Z', Z'' \in \mathcal{U}(\delta, \eta)$ be two matrices that differ only in the coordinates $z_{ij} = z_{ji}$ for $j \neq i$. Then

$$\left|\frac{P_{\Omega}(Z')}{P_{\Omega}(Z'')}\right| \leq e^{6\delta},$$

and the angle between $P_{\Omega}(Z') \neq 0$ and $P_{\Omega}(Z'') \neq 0$ does not exceed

$$2\delta \tan \frac{\theta}{2} + 5\eta.$$

Proof. It follows from (2.1) and Lemma 2.2 that

$$(3.1) |P_{\Omega}(Z)| \geq \frac{\cos(\theta/2)}{m-|\Omega|} \sum_{j \notin \Omega} \left| P_{\Omega \cup \{j\}}(Z) \right| \geq \frac{1}{(m-1)\sqrt{2}} \sum_{j \notin \Omega} \left| P_{\Omega \cup \{j\}}(Z) \right|.$$

In particular, Part 1 follows.

To prove Part 2, let us choose a branch of $\ln P_{\Omega}(Z)$ for $Z \in \mathcal{U}(\delta, \eta)$. For $0 \leq t \leq 1$, let Z(t) = tZ'' + (1-t)Z'. Then

$$\ln P_{\Omega}(Z'') - \ln P_{\Omega}(Z') = \int_0^1 \frac{d}{dt} \ln P_{\Omega}(Z(t)) dt$$
$$= \int_0^1 \sum_{j: \ j \neq i} \left(z_{ij}'' - z_{ij}' \right) \frac{\partial}{\partial z_{ij}} \ln P_{\Omega}(Z) \Big|_{Z=Z(t)} dt.$$

Using (2.2), we conclude that

$$\frac{\partial}{\partial z_{ij}} \ln P_{\Omega}(Z) = \begin{cases} 1 & \text{if } j \in \Omega, \\ P_{\Omega \cup \{j\}}(Z) / P_{\Omega}(Z) & \text{if } j \notin \Omega, \end{cases}$$

and hence

(3.2)
$$\ln P_{\Omega}(Z'') - \ln P_{\Omega}(Z') = \sum_{j \in \Omega, j \neq i} \left(z_{ij}'' - z_{ij}' \right) \\ + \int_{0}^{1} \sum_{j \notin \Omega} \left(z_{ij}'' - z_{ij}' \right) \frac{P_{\Omega \cup \{j\}}(Z(t))}{P_{\Omega}(Z(t))} dt.$$

Using (3.1), we get from (3.2) that

$$\begin{aligned} |\Re \ln P_{\Omega}(Z'') - \Re \ln P_{\Omega}(Z')| &\leq 2\delta + (m-1)\sqrt{2} \max_{j \notin \Omega} \left| z_{ij}'' - z_{ij}' \right| \\ &\leq 2\delta + 4\delta = 6\delta, \end{aligned}$$

and hence

$$\left|\frac{P_{\Omega}(Z')}{P_{\Omega}(Z'')}\right| \leq e^{6\delta},$$

as claimed.

From (2.1), for all $Z \in \mathcal{U}(\delta, \eta)$ we have that

$$\sum_{j \notin \Omega} \frac{P_{\Omega \cup \{j\}}(Z)}{P_{\Omega}(Z)} = m - |\Omega|$$

is real, while from (3.1), we conclude that

$$\sum_{j \notin \Omega} \left| \frac{P_{\Omega \cup \{j\}}(Z)}{P_{\Omega}(Z)} \right| \leq \frac{m - |\Omega|}{\cos(\theta/2)} \leq \frac{m - 1}{\cos(\theta/2)}.$$

Applying Lemma 2.1 with $u_j = P_{\Omega \cup \{j\}}(Z)/P_{\Omega}(Z)$, we conclude that

$$\sum_{j \notin \Omega} \left| \Im \frac{P_{\Omega \cup \{j\}}(Z)}{P_{\Omega}(Z)} \right| \leq (m-1) \tan \frac{\theta}{2}.$$

Therefore, from (3.2),

$$\begin{aligned} |\Im \ln P_{\Omega}(Z'') - \Im \ln P_{\Omega}(Z')| &\leq 2\eta + (m-1) \tan \frac{\theta}{2} \max_{j \notin \Omega} \left| \Re z_{ij}'' - \Re z_{ij}' \right| \\ &+ (m-1)\sqrt{2} \max_{j \notin \Omega} \left| \Im z_{ij}'' - \Im z_{ij}' \right| \\ &\leq 2\delta \tan \frac{\theta}{2} + 5\eta. \end{aligned}$$

Hence the angle between $P_{\Omega}(Z'')$ and $P_{\Omega}(Z')$ does not exceed $2\delta \tan \frac{\theta}{2} + 5\eta$, as claimed.

The second lemma shows that $P_{\Omega}(Z)$ does not change much if only the entries of Z in the *i*th row and column are changed for some $i \notin \Omega$, assuming that $n \gg m$.

LEMMA 3.2. Let us fix an $\Omega \subset \{1, \ldots, n\}$, $|\Omega| \leq m-1$. Suppose for any $i, j \notin \Omega$ and all $Z \in \mathcal{U}(\delta, \eta)$ we have $P_{\Omega \cup \{i\}}(Z) \neq 0$, $P_{\Omega \cup \{j\}}(Z) \neq 0$ and the angle between the two complex numbers does not exceed $\pi/2$ and that

$$\left|\frac{P_{\Omega\cup\{i\}}(Z)}{P_{\Omega\cup\{j\}}(Z)}\right| \leq \lambda$$

for some $\lambda \geq 1$.

In addition, suppose that if $|\Omega| \leq m-2$, then for any distinct $i, j, k \notin \Omega$ and all $Z \in \mathcal{U}(\delta, \eta)$ we have $P_{\Omega \cup \{i,j\}}(Z) \neq 0$, $P_{\Omega \cup \{i,k\}}(Z) \neq 0$, and the angle between the two complex numbers does not exceed $\pi/2$.

Let us fix an $i \notin \Omega$, and let $Z', Z'' \in \mathcal{U}(\delta, \eta)$ be two matrices that differ only in the coordinates $z_{ij} = z_{ji}$ for $j \neq i$. Then

$$\left|\frac{P_{\Omega}(Z')}{P_{\Omega}(Z'')}\right| \leq \exp\left\{\frac{10\delta\lambda m}{n-1}\right\},\,$$

and the angle between $P_{\Omega}(Z') \neq 0$ and $P_{\Omega}(Z'') \neq 0$ does not exceed

$$\frac{10\delta\lambda m}{n-1}.$$

Proof. It follows from Lemma 3.1 that $P_{\Omega}(Z) \neq 0$ for all $Z \in \mathcal{U}(\delta, \eta)$.

Arguing as in the proof of Lemma 3.1, we introduce Z(t) = tZ'' + (1-t)Z' and write

$$\ln P_{\Omega}(Z'') - \ln P_{\Omega}(Z') = \int_0^1 \sum_{j: \ j \neq i} \left(z_{ij}'' - z_{ij}' \right) \frac{\partial}{\partial z_{ij}} \ln P_{\Omega}(Z) \Big|_{Z=Z(t)} dt.$$

From (2.2), we write

(3.3)
$$\ln P_{\Omega}(Z'') - \ln P_{\Omega}(Z') = \int_{0}^{1} \sum_{j \in \Omega} \left(z_{ij}'' - z_{ij}' \right) \frac{P_{\Omega \cup \{i\}} \left(Z(t) \right)}{P_{\Omega} \left(Z(t) \right)} + \sum_{j \notin \Omega, j \neq i} \left(z_{ij}'' - z_{ij}' \right) \frac{P_{\Omega \cup \{i,j\}} \left(Z(t) \right)}{P_{\Omega} \left(Z(t) \right)} dt.$$

Suppose first that $|\Omega| \leq m-2$. From (2.1), we have

$$P_{\Omega \cup \{i\}}(Z) = \frac{1}{m - |\Omega| - 1} \sum_{j \notin \Omega, j \neq i} P_{\Omega \cup \{i, j\}}(Z).$$

Applying Lemma 2.2, we get that

(3.4)
$$\sum_{j\notin\Omega, j\neq i} \left| P_{\Omega\cup\{i,j\}}(Z) \right| \leq (m-1)\sqrt{2} \left| P_{\Omega\cup\{i\}}(Z) \right|$$

for all $Z \in \mathcal{U}(\delta, \eta)$.

Since by (2.1) we also have

$$P_{\Omega}(Z) = \frac{1}{m - |\Omega|} \sum_{j \notin \Omega} P_{\Omega \cup \{j\}}(Z)$$

applying Lemma 2.2, we conclude that

$$\sum_{j \notin \Omega} \left| P_{\Omega \cup \{j\}}(Z) \right| \leq (m - |\Omega|) \sqrt{2} \left| P_{\Omega}(Z) \right|.$$

Hence for all $i \notin \Omega$, we have

(3.5)
$$|P_{\Omega \cup \{i\}}(Z)| \leq \frac{\lambda(m-|\Omega|)\sqrt{2}}{n-|\Omega|} |P_{\Omega}(Z)| \leq \frac{\lambda m\sqrt{2}}{n} |P_{\Omega}(Z)|.$$

Combining (3.5) and (3.4), we get

(3.6)
$$\sum_{j \notin \Omega, j \neq i} \left| P_{\Omega \cup \{i,j\}}(Z) \right| \leq \frac{2\lambda m(m-1)}{n} \left| P_{\Omega}(Z) \right|.$$

Combining (3.3), (3.4), (3.5), and (3.6), we get

$$\begin{aligned} |\ln P_{\Omega}(Z'') - \ln P_{\Omega}(Z')| &\leq \frac{2\sqrt{2}\delta}{m-1} \cdot \frac{\lambda |\Omega|(m-|\Omega|)\sqrt{2}}{n-|\Omega|} + \frac{2\sqrt{2}\delta}{m-1} \cdot \frac{2\lambda m(m-1)}{n} \\ &\leq \frac{4\delta\lambda m}{n-1} + \frac{4\sqrt{2}\delta\lambda m}{n} \leq \frac{10\delta\lambda m}{n-1}. \end{aligned}$$

If $|\Omega| = m - 1$, then from (3.3) and (3.5), we get

$$\left|\ln P_{\Omega}(Z'') - \ln P_{\Omega}(Z')\right| \leq \frac{2\sqrt{2}\delta}{m-1} \cdot \frac{\lambda m\sqrt{2}}{n} \leq \frac{4\delta\lambda m}{n-1},$$

which concludes the proof.

Now we are ready to prove Theorem 1.3.

Proof of Theorem 1.3. Given $0 < \delta < 1$, we choose $0 < \theta < \pi/2$ so that

$$2\delta \tan \frac{\theta}{2} < \theta.$$

We then choose $\eta > 0$ such that

$$2\delta \tan \frac{\theta}{2} + 5\eta < \theta$$

We choose

$$\lambda > e^{6\delta}$$

and choose $\omega > 1$ so that

$$2\delta \tan \frac{\theta}{2} + 5\eta + \frac{10\delta\lambda m}{n-1} \leq \theta \quad \text{and} \quad \exp\left\{6\delta + \frac{10\delta\lambda m}{n-1}\right\} \leq \lambda$$

whenever $n \geq \omega m$.

Suppose that $n \ge \omega m$. We prove by descending induction on $r = m, m - 1, \ldots, 1$ that if $\Omega_1, \Omega_2 \in \{1, \ldots, n\}$ are two sets such that $|\Omega_1| = |\Omega_2| = r$ and $|\Omega_1 \Delta \Omega_2| = 2$, then for all $Z \in \mathcal{U}(\delta, \eta)$ we have $P_{\Omega_1}(Z) \ne 0$, $P_{\Omega_2}(Z) \ne 0$, the angle between $P_{\Omega_1}(Z)$ and $P_{\Omega_2}(Z)$ does not exceed θ while the ratio of $|P_{\Omega_1}(Z)|$ and $|P_{\Omega_2}(Z)|$ does not exceed λ .

Assume that r = m. Without loss of generality, we assume that $\Omega_1 = \Omega \cup \{1\}$ and $\Omega_2 = \Omega \cup \{2\}$ for some $\Omega \subset \{3, \ldots, n\}$ such that $|\Omega| = m - 1$. We have

$$P_{\Omega_1}(Z) = \exp\left\{\sum_{\{i,j\}\subset\Omega} z_{ij}\right\} \exp\left\{\sum_{i\in\Omega} z_{1i}\right\} \text{ and}$$
$$P_{\Omega_2}(Z) = \exp\left\{\sum_{\{i,j\}\subset\Omega} z_{ij}\right\} \exp\left\{\sum_{i\in\Omega} z_{2i}\right\}.$$

Clearly, $P_{\Omega_1}(Z) \neq 0$, $P_{\Omega_2}(Z) \neq 0$, the angle between $P_{\Omega_1}(Z)$ and $P_{\Omega_2}(Z)$ does not exceed $2\eta \leq \theta$ while the ratio of $|P_{\Omega_1}(Z)|$ and $|P_{\Omega_2}(Z)|$ does not exceed $e^{2\delta} \leq \lambda$.

Suppose now that the statements hold for all subsets $\Omega \subset \{1, \ldots, n\}$ of cardinality at least r + 1 for some $r \leq m - 1$, and let $\Omega_1, \Omega_2 \subset \{1, \ldots, n\}$ be two subsets of cardinality $r \geq 1$ such that $|\Omega_1 \Delta \Omega_2| = 2$. Again, without loss of generality, we assume that $\Omega_1 = \Omega \cup \{1\}$ and $\Omega_2 = \Omega \cup \{2\}$ for some $\Omega \subset \{3, \ldots, n\}$ such that $|\Omega| = r - 1$. Then we observe that $P_{\Omega_2}(Z) = P_{\Omega_1}(Z')$, where

$$z'_{1i} = z'_{11} = z_{2i} = z_{i2}$$
 and $z'_{2i} = z'_{12} = z_{1i} = z_{i1}$ for $i \neq 1, 2,$

while all other entries of Z and Z' coincide. Applying Lemma 3.1 and Lemma 3.2 and the induction hypothesis to sets $\Omega_1 \cup \{j\}$ for $j \notin \Omega_1$ and $\Omega_1 \cup \{j, k\}$ for $j, k \notin \Omega_1$, we conclude that the angle between $P_{\Omega_1}(Z) \neq 0$ and $P_{\Omega_2}(Z) \neq 0$ does not exceed

$$2\delta \tan \frac{\theta}{2} + 5\eta + \frac{10\delta\lambda m}{n-1} \leq \theta,$$

while the ratio of $|P_{\Omega_1}(Z)|$ and $|P_{\Omega_2}(Z)|$ does not exceed

$$\exp\left\{6\delta + \frac{10\delta\lambda m}{n-1}\right\} \le \lambda.$$

This proves that $P_{\{i\}}(Z) \neq 0$ for all $i \in \{1, \ldots, n\}$ and all $Z \in \mathcal{U}(\delta, \eta)$ and that the angle between $P_{\{i\}}(Z) \neq 0$ and $P_{\{j\}}(Z) \neq 0$ does not exceed θ for all $i, j \in \{1, \ldots, n\}$. From (2.1) we conclude that $P_m(Z) = P_{\emptyset}(Z) \neq 0$ for all $Z \in \mathcal{U}(\delta, \eta)$.

4. Computing the partition function. Here we show how to compute the density partition function $den_m(G; \gamma)$. First, we make a change of coordinates to convert the partition function $P_m(Z)$ of section 1.2 into a multivariate polynomial.

4.1. A polynomial version of $P_m(Z)$. For an $n \times n$ complex symmetric matrix $W = (w_{ij})$ with zero diagonal, we define

$$p_m(W) = \binom{n}{m}^{-1} \sum_{\substack{S \subset \{1, \dots, n\} \\ |S| = m}} \prod_{\substack{\{i, j\} \subset S \\ i \neq j}} (1 + w_{ij}).$$

Hence $p_m(W)$ is a polynomial of degree $\binom{m}{2}$ in the entries w_{ij} and, assuming that $|w_{ij}| < 1$ for all i, j, we can write

$$p_m(W) = \binom{n}{m}^{-1} P_m(Z), \quad \text{where} \quad Z = (z_{ij}) \quad \text{and} \quad z_{ij} = \ln\left(1 + w_{ij}\right)$$

(we choose the standard branch of the logarithm in the right half-plane of \mathbb{C}). Theorem 1.3 implies that for every $0 < \delta < 1$ there is $\eta = \eta(\delta) > 0$ and $\omega = \omega(\delta) > 1$ such that

(4.1)

$$p_m(W) \neq 0 \quad \text{whenever} \quad |\Re \ln (1 + w_{ij})| \leq \frac{\delta}{m-1},$$

$$|\Im \ln (1 + w_{ij})| \leq \frac{\eta}{m-1}, \quad \text{and}$$

$$n \geq \omega m.$$

To compute den_m(G; γ) for a given $0 < \gamma < 1$ and a given graph G = (V, E), we define

(4.2)
$$w_{ij} = \begin{cases} \exp\left\{\frac{\gamma}{m-1}\right\} - 1 & \text{if } \{i,j\} \in E, \\ \exp\left\{-\frac{\gamma}{m-1}\right\} - 1 & \text{if } \{i,j\} \notin E. \end{cases}$$

Then, by (1.6), we have

(4.3)
$$\operatorname{den}_m(G;\gamma) = \exp\left\{\frac{\gamma m}{2}\right\} p_m(W).$$

The interpolation method is based on the following simple lemma.

LEMMA 4.1. Let $g: \mathbb{C} \longrightarrow \mathbb{C}$ be a univariate polynomial, and suppose that $g(z) \neq 0$ provided $|z| < \beta$, where $\beta > 1$ is some real number. Let us choose a branch of $f(z) = \ln g(z)$ in the disc $|z| < \beta$, and let

$$T_r(z) = f(0) + \sum_{k=1}^r \frac{f^{(k)}(0)}{k!} z^k$$

be the Taylor polynomial of f of degree r computed at z = 0. Then

$$|f(1) - T_r(1)| \leq \frac{\deg g}{\beta^r(\beta - 1)(r+1)}.$$

Proof. This is Lemma 2.2.1 of [4]; see also Lemma 1.1 of [3].

The gist of Lemma 4.1 is that to approximate f(1) within an additive error ϵ ; it suffices to compute the Taylor polynomial of f(z) at 0 of degree $r = O_{\beta} (\ln \deg g - \ln \epsilon)$, where the implicit constant in the "O" notation depends on β alone. We would like to apply Lemma 4.1 to the univariate polynomial

(4.4)
$$h(z) = \binom{n}{m}^{-1} \sum_{\substack{S \subset \{1, \dots, n\} \\ |S| = m}} \prod_{\substack{\{i, j\} \subset S \\ i \neq j}} (1 + zw_{ij}),$$

where w_{ij} are defined by (4.2). Indeed, the value we are ultimately interested is $h(1) = p_m(W)$. However, Lemma 4.1 requires that $h(z) \neq 0$ in a disc of some radius $\beta > 1$, whereas (4.1) only guarantees that $h(z) \neq 0$ for z in a neighborhood of the interval $[0,1] \subset \mathbb{C}$. To remedy this, we compose h with a polynomial $\phi : \mathbb{C} \longrightarrow \mathbb{C}$ such that $\phi(0) = 0$, $\phi(1) = 1$ and ϕ maps the disc $|z| < \beta$ for some $\beta > 1$ inside the prescribed neighborhood of $[0,1] \subset \mathbb{C}$. We then apply Lemma 4.1 to the composition $g(z) = h((\phi(z)))$. The following lemma provides an explicit construction of ϕ .

LEMMA 4.2. For $0 < \rho < 1$, we define

$$\begin{aligned} \alpha &= \alpha(\rho) = 1 - e^{-\frac{1}{\rho}}, \quad \beta = \beta(\rho) = \frac{1 - e^{-1 - \frac{1}{\rho}}}{1 - e^{-\frac{1}{\rho}}} > 1, \\ N &= N(\rho) = \left\lfloor \left(1 + \frac{1}{\rho}\right) e^{1 + \frac{1}{\rho}} \right\rfloor, \quad \sigma = \sigma(\rho) = \sum_{k=1}^{N} \frac{\alpha^{k}}{k}, \quad and \\ \phi(z) &= \phi_{\rho}(z) = \frac{1}{\sigma} \sum_{k=1}^{N} \frac{(\alpha z)^{k}}{k}. \end{aligned}$$

Then $\phi : \mathbb{C} \longrightarrow \mathbb{C}$ is a polynomial of degree N such that $\phi(0) = 0, \ \phi(1) = 1$,

$$-
ho \ \le \ \Re \, \phi(z) \ \le \ 1 + 2
ho, \quad and \quad |\Im \, \phi(z)| \ \le \ 2
ho,$$

provided $|z| \leq \beta$.

Proof. This is Lemma 2.2.3 of [4].

Lemma 4.1 also requires the derivatives $f^{(k)}(0)$ of $f(z) = \ln g(z)$ at z = 0. Those, however, can be easily computed from the derivatives $g^{(k)}(0)$, as described in section 2.2.2 of [4]; see also section 2.1 of [3]. We briefly sketch how.

4.2. Computing derivatives. Suppose that $f(z) = \ln g(z)$ as in Lemma 4.1. Then

$$f'(z) = \frac{g'(z)}{g(z)}$$
 and $g'(z) = f'(z)g(z)$

Differentiating the product k-1 times, we obtain

(4.5)
$$g^{(k)}(0) = \sum_{j=0}^{k-1} \binom{k-1}{j} f^{(k-j)}(0) g^{(j)}(0) \quad \text{for} \quad k = 1, \dots, r.$$

We interpret (4.5) as a system of linear equations in variables $f^{(k)}(0)$ for k = 1, ..., rwith coefficients $g^{(k)}(0)$ for k = 0, ..., r. This is a triangular system of linear equations

with nonzero entries $g^{(0)}(0) = g(0)$ on the diagonal, that can be solved in $O(r^2)$ time, provided the values of $g^{(k)}(0)$ are known.

To supply the last ingredient of the algorithm, we show how to compute $h^{(k)}(0)$ for k = 0, ..., r, where h is the polynomial defined by (4.4). This is also done in [3], but we reproduce it here for completeness.

We have

$$h^{(k)}(0) = \binom{n}{m}^{-1} \sum_{\substack{S \subset \{1, \dots, n\} \\ |S| = m}} \sum_{\substack{\{i_1, j_1\}, \dots, \{i_k, j_k\} \subset S}} w_{i_1 j_1} \dots w_{i_k j_k},$$

where the inner sum is taken over all ordered collections of distinct unordered pairs $\{i_1, j_1\}, \ldots, \{i_k, j_k\} \subset S$. For such a collection, say I, let $\nu(I)$ be the number of distinct vertices among $i_1, j_1, \ldots, i_k, j_k$. Then there are exactly $\binom{n-\nu(I)}{m-\nu(I)}$ different m-subsets S containing the edges from I, and we can rewrite the above sum as

(4.6)
$$h^{(k)}(0) = \binom{n}{m}^{-1} \sum_{I = (\{i_1, j_1\}, \dots, \{i_k, j_k\})} \binom{n - \nu(I)}{m - \nu(I)} w_{i_1 j_1} \dots w_{i_k j_k}$$

where the sum is taken over all ordered collections of k unordered pairs $\{i_s, j_s\}$. It is clear now that $h^{(k)}(0)$ can be computed in $n^{O(k)}$ time by the exhaustive enumeration of all possible collections of k pairs.

In section 5 we present faster formulas for computing $h^{(2)}(0)$ and $h^{(3)}(0)$ that we used for our numerical experiments.

4.3. The algorithm. Let us fix $0 < \gamma < 1$. Below we summarize the algorithm for computing den_m(G; γ) within relative error $0 < \epsilon < 1$, by which we understand computing $\ln \operatorname{den}_m(G; \gamma)$ within additive error ϵ . We assume that $m \ge 4$ and that $n \ge \omega m$ for some $\omega = \omega(\gamma) > 1$, to be specified below.

Given a graph G = (V, E) with set $V = \{1, \ldots, n\}$ of vertices, and an integer $m \leq n$, we compute the $n \times n$ symmetric matrix $W = (w_{ij})$ by (4.2). Since $m \geq 4$, we have $|w_{ij}| \leq 0.4$ for all i, j.

Our goal is to compute $p_m(W) = h(1)$, where h is the univariate polynomial defined by (4.4). We note that deg $h = \binom{m}{2}$.

Let us choose $1 > \delta > \gamma$, and let $\eta = \eta(\delta) > 0$ and $\omega = \omega(\delta) > 1$ be the numbers of Theorem 1.3 and in (4.1). We find $\rho = \rho(\delta) > 0$ such that

$$|\Re \ln (1+zw_{ij})| \leq rac{\delta}{m-1}$$
 and $|\Im \ln (1+zw_{ij})| \leq rac{\eta}{m-1}$

as long as

$$(4.7) -\rho \leq \Re z \leq 1+\rho \text{ and } |\Im z| \leq \rho.$$

Indeed, if $z \in [0, 1]$, then

$$-\frac{\gamma}{m-1} \leq \ln\left(1+zw_{ij}\right) \leq \frac{\gamma}{m-1},$$

and for $|z| \leq 2$, we have

$$\left|\frac{d}{dz}\ln\left(1+zw_{ij}\right)\right| = \left|\frac{w_{ij}}{1+zw_{ij}}\right| \leq \frac{10}{m-1},$$

so the desired ρ can indeed be found.

It follows by (4.1) that $h(z) \neq 0$ as long as $n \geq \omega m$ and (4.7) holds.

Using Lemma 4.2, we construct a polynomial $\phi : \mathbb{C} \longrightarrow \mathbb{C}$ of some degree $N = N(\rho) = N(\delta)$ such that $\phi(0) = 0$, $\phi(1) = 1$ and

$$-\rho \leq \Re \phi(z) \leq 1+\rho \text{ and } |\Im \phi(z)| \leq \rho$$

as long as $|z| \leq \beta$ for some $\beta = \beta(\rho) = \beta(\delta) > 1$. We define

$$g(z) = h(\phi(z)),$$

and our goal is to compute $g(1) = h(\phi(1))$. We note that

$$\deg g \leq N \deg h = N\binom{m}{2}.$$

We choose a branch of $f(z) = \ln g(z)$ for z satisfying (4.7).

Using Lemma 4.1, we find an integer $r = O_{\rho} (\ln m - \ln \epsilon) = O_{\delta} (\ln m - \ln \epsilon)$ such that

$$|T_r(1) - f(1)| \leq \epsilon,$$

where $T_r(z)$ is the Taylor polynomial of f(z) of degree r, computed at z = 0. The implicit constant in the "O" notation depends only on ρ , which in turn depends only on δ . Hence our goal is to compute $T_r(1)$, for which we need to compute $f^{(k)}(0)$ for $k = 1, \ldots, r$. As in section 4.2, we reduce it in $O(r^2)$ time to computing $g^{(k)}(0)$ for $k = 1, \ldots, r$. Note that

$$g(0) = h(\phi(0)) = h(0) = 1.$$

Let $\phi_r(z)$ be the truncation of the polynomial $\phi(z)$ obtained by discarding all monomials of degree higher than r. Similarly, let $h_r(z)$ be the truncation of the polynomial h(z), obtained by discarding all monomial of degree higher than r. We compute $h_r(z)$ as in section 4.2 in $n^{O(r)}$ time. Finally, we compute the truncation of the composition $h_r(\phi_r(z))$. A fast (polynomial in r) way to do it, is to use Horner's method: assuming that

$$h_r(z) = \sum_{k=0}^r b_k z^k,$$

we successively compute

$$b_r \phi_r(z) + b_{r-1}, \quad (b_r \phi_r(z) + b_{r-1}) \phi_r(z) + b_{r-2}, ((b_r \phi_r(z) + b_{r-1}) \phi_r(z) + b_{r-2}) \phi_r(z) + b_{r-3}, \dots$$

discarding on the way all monomials of degree higher than r. In the end, we have computed $g^{(k)}(0)$ for k = 0, ..., r, and hence $f^{(k)}(0)$ for k = 0, ..., r, and hence $T_m(1)$ approximating $f(1) = \ln h(1)$ within additive error ϵ . From (4.3), we compute

$$\operatorname{den}_m(G;\gamma) = \exp\left\{\frac{\gamma m}{2}\right\} h(1)$$

within relative error $\epsilon > 0$.

5. Remarks on the practical implementation. We implemented a *much* simplified version of the algorithm. Given a graph G = (V, E) with set $V = \{1, \ldots, n\}$ of vertices and an integer $2 \le m \le n$, we define the $n \times n$ matrix $= (w_{ij})$ by

$$w_{ij} = \begin{cases} \alpha & \text{if } \{i, j\} \in E \\ -\alpha & \text{if } \{i, j\} \notin E, \end{cases}$$

where $0 < \alpha < 1$ is a parameter.

We consider the polynomial h(z) defined by (4.4) and let $f(z) = \ln h(z)$. Our goal is to approximate $f(1) = \ln h(1)$, and hence

$$h(1) = \sum_{\substack{S \subset \{1, \dots, n\} \\ |S| = m}} (1+\alpha)^{\binom{m}{2}\sigma(S)} (1-\alpha)^{\binom{m}{2}(1-\sigma(S))}$$
$$= (1-\alpha)^{\binom{m}{2}} \operatorname{den}_m(G;\gamma), \quad \text{where} \quad \gamma = \frac{m-1}{2} \ln \frac{1+\alpha}{1-\alpha}$$

We approximate f(1) by the degree r Taylor polynomial of f(z) computed at z = 0. The results of [3] suggest that for $\alpha = O(1/m)$, we should get a reasonable approximation if we use $r \sim \ln m$. The results of our numerical experiments suggest that we get reasonable approximations if we use $\alpha = \Omega(1)$ and r = 2 or r = 3. In short, on the examples we tested, the quality of approximation was more consistent with the quality of the Taylor polynomial approximation of $\ln(1 \pm \alpha)$.

More precisely, we ran the algorithm typically with parameters n = 50, 100 and m = 10, although occasionally we chose n as large as n = 300. For the parameters n = 50 and m = 10 we were able to compare our approximation with the exact value. Typically, choosing $\alpha = 0.5$ or lower produced an approximation of f(1) within 1% accuracy. For $\alpha = 0.7$, the accuracy went down to 10% - 20% and for $\alpha > 0.7$ the approximation was not accurate. For higher values of n, where the exact value of f(1) was unavailable, we compared the approximations obtained for r = 2 and r = 3. If the approximations were close to each other, we considered it as an indication that they are also close to the true value of f(1). Again, we observed that up to $\alpha = 0.5$, the approximations agreed but were beginning to essentially differ at $\alpha = 0.7$ and higher. For the graphs, we used the Erdős–Rényi models G(n, 0.5), G(n, 0.4), those graphs with planted cliques of size m, and occasionally manually constructed "random-looking" graphs.

We provide below the explicit formulas for the approximations up to degree 3, in case the reader will be interested to do some numerical experiments. We interpret w_{ij} as weights on the edges of a complete graph with *n* vertices. Borrowing an idea from [13], we express the derivatives $f^{(k)}(0)$ in terms of various sums associated with *connected* subgraphs, since it improves the computational complexity of the algorithm. We remark, however, that it looks unlikely that the methods of [13] can be pushed to improve the complexity of our algorithm in the general situation from quasi-polynomial to genuinely polynomial, since we work with graphs of unbounded degrees.

It is convenient to introduce the following sums:

$$A_1 = \sum_{\{i,j\}} w_{ij},$$

where the sum is taken over all unordered pairs $\{i, j\}$ of distinct indices;

$$B_1 = \sum_{\{i,j\}} w_{ij}^2, \quad B_2 = \sum_{j,\{i,k\}} w_{ij} w_{jk},$$

where in the formula for B_1 the sum is taken oven all unordered pairs $\{i, j\}$ of distinct indices and in B_2 the sum is taken over all pairs consisting of an index j and an unordered pair $\{i, k\}$, so that all three indices are distinct; and

$$\begin{split} C_1 &= \sum_{\{i,j\}} w_{ij}^3, \quad C_2 = \sum_{(i,j,k)} w_{ij}^2 w_{jk}, \quad C_3 = \sum_{\{i,j,k\}} w_{ij} w_{jk} w_{ki}, \\ C_4 &= \sum_{(i,j,k,l)} w_{ij} w_{jk} w_{kl}, \quad C_5 = \sum_{\{j,k,l\},i} w_{il} w_{ij} w_{ik}, \end{split}$$

where in C_1 the sum is taken over all unordered pairs $\{i, j\}$ of distinct indices, in C_2 the sum is taken over all ordered triples (i, j, k) of distinct indices, in C_3 the sum is taken over all unordered triples of distinct integers, in C_4 the sum is taken over all ordered 4-tuples (i, j, k, l) of distinct indices, and in C_5 the sum is taken over all pairs consisting of an index i and an unordered triple $\{j, k, l\}$ so that all four indices $\{i, j, k, l\}$ are distinct.

5.1. First-order approximation. Clearly, h(0) = 1. From (4.6), we have

$$h'(0) = \binom{n}{m}^{-1} \binom{n-2}{m-2} \sum_{\{i,j\} \subset \{1,\dots,n\}} w_{ij} = \frac{m(m-1)}{n(n-1)} A_1.$$

Since $f(0) = \ln h(0) = 0$ and f'(0) = h'(0)/h(0) = h'(0), we obtain the first order approximation

$$f(1) \approx h'(0),$$

where h'(0) is defined as above. The complexity of computing the first order approximation is $O(n^2)$.

5.2. Second-order approximation. From (4.6), we have

$$h''(0) = \binom{n}{m}^{-1} \sum_{I = (\{i_1, j_1\}, \{i_2, j_2\})} \binom{n - \nu(I)}{m - \nu(I)} w_{i_1 j_1} w_{i_2 j_2}.$$

Here $\nu(I) = 4$ if the pairs $\{i_1, j_1\}$ and $\{i_2, j_2\}$ are pairwise disjoint and $\nu(I) = 3$ if they share exactly one index. Hence we can write

$$h''(0) = \binom{n}{m}^{-1} \left(2\binom{n-3}{m-3} B_2 + \binom{n-4}{m-4} \left(A_1^2 - 2B_2 - B_1 \right) \right)$$

= $2\frac{m(m-1)(m-2)}{n(n-1)(n-2)} B_2 + \frac{m(m-1)(m-2)(m-3)}{n(n-1)(n-2)(n-3)} \left(A_1^2 - 2B_2 - B_1 \right)$

Since

$$f''(0) = h''(0) - (h'(0))^2$$

we obtain the second order approximation:

$$f(1) \approx f'(0) + \frac{1}{2}f''(0) = h'(0) - \frac{1}{2}(h'(0))^2 + \frac{1}{2}h''(0),$$

where h'(0) and h''(0) are defined as above. The complexity of computing the second order approximation is $O(n^3)$.

5.3. Third-order approximation. From (4.6), one can deduce that

$$\begin{split} h^{\prime\prime\prime}(0) &= 6\frac{m(m-1)(m-2)}{n(n-1)(n-2)}C_3 + \frac{m(m-1)(m-2)(m-3)}{n(n-1)(n-2)(n-3)}\left(6C_5 + 3C_4\right) \\ &+ 6\frac{m(m-1)(m-2)(m-3)(m-4)}{n(n-1)(n-2)(n-3)(n-4)}\left(A_1B_2 - 3C_5 - 3C_3 - C_4 - C_2\right) \\ &+ \frac{m(m-1)(m-2)(m-3)(m-4)(m-5)}{n(n-1)(n-2)(n-3)(n-4)(n-5)}\left(A_1^3 + 12C_3 - 6A_1B_2 + 12C_5 + 3C_4 + 6C_2 - 3A_1B_1 + 2C_1\right). \end{split}$$

Since we have

$$f'''(0) = h'''(0) - 2f''(0)h'(0) - f'(0)h''(0) = 2(h'(0))^3 - 3h'(0)h''(0) + h'''(0),$$

we obtain the third order approximation approximation

$$f(1) \approx f'(0) + \frac{1}{2}f''(0) + \frac{1}{6}f'''(0)$$

= $h'(0) - \frac{1}{2}(h'(0))^2 + \frac{1}{2}h''(0) + \frac{1}{3}(h'(0))^3 - \frac{1}{2}h'(0)h''(0) + \frac{1}{6}h'''(0)$

The complexity of computing the third order approximation is $O(n^4)$.

6. Proof of Theorem 1.4 and concluding remarks. We got the idea of the proof from [9], where a similar question about complex zeros of the permanents of matrices with independent random entries was treated.

Proof of Theorem 1.4. Applying Jensen's formula (see, for example, section 5.3 of [1]), we obtain

(6.1)
$$\ln|h_W(0)| = \sum_{s=1}^N \ln \frac{|a_{s,W}|}{r} + \frac{1}{2\pi} \int_0^{2\pi} \ln |h_W(re^{i\theta})| \ d\theta,$$

where $a_{s,W}$, s = 1, ..., N are the roots of the polynomial $h_W(z)$ in the disc |z| < r, and we assume that $h_W(z)$ has no zeros on the circle |z| = r (since there are only finitely many values of r with roots on the circle |z| = r, this assumption is not restrictive). We have

$$\ln|h_W(0)| = 0,$$

and furthermore, applying Jensen's inequality, we bound:

(6.2)
$$\frac{1}{2\pi} \int_0^{2\pi} \ln \left| h_W \left(r e^{i\theta} \right) \right| \ d\theta = \frac{1}{2} \cdot \frac{1}{2\pi} \int_0^{2\pi} \ln \left| h_W \left(r e^{i\theta} \right) \right|^2 \ d\theta$$
$$\leq \frac{1}{2} \ln \left(\frac{1}{2\pi} \int_0^{2\pi} \left| h_W \left(r e^{i\theta} \right) \right|^2 \ d\theta \right).$$

For a fixed $\theta \in [0, 2\pi]$, we compute the expectation

$$\mathbf{E} \left| h_{W} \left(re^{i\theta} \right) \right|^{2} = {\binom{n}{m}}^{-2} \sum_{\substack{S_{1}, S_{2} \subset \{1, \dots, n\} \\ |S_{1}| = |S_{2}| = m}} \mathbf{E} \left(\prod_{\{j, k\} \subset S_{1}} \left(1 + re^{i\theta} w_{jk} \right) \right)$$
$$\times \prod_{\{j, k\} \subset S_{2}} \left(1 + re^{-i\theta} w_{jk} \right) \right)$$
$$= {\binom{n}{m}}^{-2} \sum_{\substack{S_{1}, S_{2} \subset \{1, \dots, n\} \\ |S_{1}| = |S_{2}| = m}} \left(1 + r^{2} \right)^{\binom{|S_{1} \cap S_{2}|}{2}}.$$

A subset $S \subset \{1, \ldots, n\}$ of cardinality $l = |S| \leq m$ can be represented as the intersection $S = S_1 \cap S_2$ of *m*-subsets S_1, S_2 in $\binom{n-l}{m-l}\binom{n-m}{m-l}$ ways. Hence

(6.3)
$$\mathbf{E} \left| h_W \left(r e^{i\theta} \right) \right|^2 = {\binom{n}{m}}^{-2} \sum_{l=0}^m {\binom{n}{l}} {\binom{n-l}{m-l}} {\binom{n-m}{m-l}} \left(1 + r^2 \right)^{\binom{l}{2}}.$$

To bound (6.3), we consider the ratio of the (l + 1)st term to the *l*th term:

$$\frac{n-l}{l+1} \cdot \frac{m-l}{n-l} \cdot \frac{m-l}{n-2m+l+1} \cdot \left(1+r^2\right)^l = \frac{(m-l)^2 \left(1+r^2\right)^l}{(l+1)(n-2m+l+1)} \\ \leq \frac{m^2(1+r^2)^m}{n-2m+1}.$$

In particular, if

(6.4)
$$n \ge 2m^2(1+r^2)^m + 2m,$$

the ratio does not exceed 1/2, and hence we can bound the sum (6.3) by

$$\mathbf{E} \left| h_W \left(r e^{i\theta} \right) \right|^2 \leq 2 \binom{n}{m}^{-2} \binom{n}{m} \binom{n-m}{m} \leq 2.$$

Integrating over θ , we conclude that if (6.4) holds, then

$$\mathbf{E}\left(\frac{1}{2\pi}\int_{0}^{2\pi}\left|h_{W}\left(re^{i\theta}\right)\right| \ d\theta\right) \leq 2.$$

By the Markov inequality, for any $\tau \geq 1$, we get

$$\mathbf{P}\left(\frac{1}{2\pi}\int_{0}^{2\pi}\left|h_{W}\left(re^{i\theta}\right) d\theta\right| \geq 2\tau\right) \leq \frac{1}{\tau}.$$

Consequently, from (6.1) and (6.2), we have

$$\mathbf{P}\left(\sum_{s=1}^{N} \ln \frac{|a_{s,W}|}{r} \leq -\frac{1}{2} \ln 2\tau\right) \leq \frac{1}{\tau},$$

and the proof follows.

An anonymous referee asked what happens if G is a random graph G(n, 0.5) with a planted *m*-clique. The most interesting asymptotic regime is when $m^2 \ll n \leq m^{O(1)}$ and *m* grows; see [2] for results and references. Here we are interested in a polynomial time algorithm which, with high probability, tells G from G(n, 0.5). A

quasi-polynomial time algorithm is readily available (by an exhaustive search for a clique of size at least $3 \log_2 n$, say). Our proof of Theorem 1.4 does not seem to extend to random graphs with a planted clique. We note, however, that if the radius of zero-free region is roughly the same $r = \Omega(1/\sqrt{m})$ as in Theorem 1.4 or even weaker, $r = \Omega(m^{-1+\epsilon})$ for some $\epsilon > 0$, we do obtain a desired polynomial time algorithm. Indeed, in the latter case, we can choose $\gamma = m^{\epsilon'}$ with some $0 < \epsilon' < \epsilon$. If G is a graph with a planted m-clique, we have

$$\operatorname{den}_m(G;\gamma) \geq \exp\left\{m^{1+\epsilon'} - O(m\ln m)\right\};$$

cf. (1.2). If G is a random graph G(n, 0.5), our proof Theorem 1.4 implies that

$$\operatorname{den}_m(G;\gamma) \leq \exp\left\{\frac{m^{1+\epsilon'}}{2} + O(1)\right\}$$

with high probability; cf. (1.7). Note that by choosing $\epsilon' < \epsilon$, we choose γ sufficiently "deep" inside the purported zero-free region, and hence we can get a genuinely polynomial, as opposed to a quasi-polynomial, algorithm by computing a constant, as opposed to logarithmic, number of terms in the Taylor polynomial approximation; cf. Lemma 4.1.

REFERENCES

- L. V. AHLFORS, Complex Analysis. An introduction to the theory of analytic functions of one complex variable, 3rd ed., McGraw-Hill Book Co., New York, 1978.
- [2] M. ALON, M. KRIVELEVICH, AND B. SUDAKOV, Finding a large hidden clique in a random graph, Random Structures Algorithms, 13 (1998), pp. 457–466.
- [3] A. BARVINOK, Computing the partition function for cliques in a graph, Theory Comput., 11 (2015), pp. 339–355, https://doi.org/10.4086/toc.2015.v011a013.
- [4] A. BARVINOK, Combinatorics and Complexity of Partition Functions, Springer, Cham, 2016.
- [5] A. BHASKARA, Finding dense structures in graphs and matrices, Ph.D. dissertation, Princeton University, 2012, https://www.cs.utah.edu/~bhaskara/files/thesis.pdf.
- [6] A. BHASKARA, M. CHARIKAR, E. CHLAMTAC, U. FEIGE, AND A. VIJAYARAGHAVAN, Detecting high log-densities – an o(n^{1/4}) approximation for densest k-subgraph, in Proceedings of the 2010 ACM International Symposium on Theory of Computing, New York, 2010, ACM, pp. 201–210.
- [7] A. BHASKARA, M. CHARIKAR, V. GURUSWAMI, A. VIJAYARAGHAVAN, AND Y. ZHOU, Polynomial integrality gaps for strong sdp relaxations of densest k-subgraph, in Proceedings of the Twenty-Third Annual ACM-SIAM Symposium on Discrete Algorithms, New York, 2012, ACM, pp. 388–405.
- [8] B. BOLLOBÁS, Combinatorics: Set systems, Hypergraphs, Families of Vectors and Combinatorial Probability, Cambridge University Press, Cambridge, 1986.
- [9] L. ELDAR AND S. MEHRABAN, Approximating the permanent of a random matrix with vanishing mean, in Proceedings of the 2018 IEEE 59th Annual Symposium on Foundations of Computer Science (FOCS 2018), Los Alamitos, CA, 2018, IEEE, pp. 23–34.
- [10] A. FRIEZE AND R. KANNAN, Quick approximation to matrices and applications, Combinatorica, 19 (1999), pp. 175–220.
- [11] O. GOLDREICH, Introduction to Property Testing, Cambridge University Press, Cambridge, 2017.
- [12] J. HÅSTAD, Clique is hard to approximate within $n^{1-\epsilon}$, Acta Mathematica, 182 (1999), pp. 105–142.
- [13] V. PATEL AND G. REGTS, Deterministic polynomial-time approximation algorithms for partition functions and graph polynomials, SIAM J. Comput., 46 (2017), pp. 1893–1919.
- [14] C. TERRELL, Mathematical Statistics: A Unified Introduction, Springer-Verlag, New York, 1999.
- [15] D. ZUCKERMAN, Linear degree extractors and the inapproximability of max clique and chromatic number, Theory Comput., 3 (2007), pp. 103–128, https://doi.org/10.4086/toc.2007. v003a006.