

# TESTING FOR DENSE SUBSETS IN A GRAPH VIA THE PARTITION FUNCTION

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ABSTRACT. For a set  $S$  of vertices of a graph  $G$ , we define its density  $0 \leq \sigma(S) \leq 1$  as the ratio of the number of edges of  $G$  spanned by the vertices of  $S$  to  $\binom{|S|}{2}$ . We show that, given a graph  $G$  with  $n$  vertices and an integer  $m \ll n$ , the partition function  $\sum_S \exp\{\gamma m \sigma(S)\}$ , where the sum is taken over all  $m$ -subsets  $S$  of vertices and  $0 < \gamma < 1$  is fixed in advance, can be approximated within relative error  $0 < \epsilon < 1$  in quasi-polynomial  $n^{O(\ln m - \ln \epsilon)}$  time. We discuss numerical experiments and observe that for the random graph  $G(n, 1/2)$  one can afford a much larger  $\gamma$ , provided the ratio  $n/m$  is sufficiently large.

## 1. INTRODUCTION AND MAIN RESULTS

Let  $G = (V, E)$  be an undirected graph, without loops or multiple edges. For a non-empty subset  $S \subset V$  of vertices, we define the *density*  $\sigma(S)$  as the fraction of the pairs of vertices of  $S$  that span an edge of  $G$ :

$$\sigma(S) = \frac{\left| \binom{S}{2} \cap E \right|}{\binom{|S|}{2}},$$

where  $\binom{S}{2}$  is the set of all unordered pairs of vertices from  $S$ . Hence  $0 \leq \sigma(S) \leq 1$  for all subsets,  $\sigma(S) = 0$  if  $S$  is an *independent set* and  $\sigma(S) = 1$  if  $S$  is a *clique*.

We are interested in the following general problem: given a graph  $G = (V, E)$  with  $|V| = n$  vertices and an integer  $m \leq n$ , estimate the highest density of an  $m$ -subset  $S \subset V$ . This is, of course, a hard problem: for example, testing whether a given graph contains a clique of a given size, or even estimating the size of the largest clique within a factor of  $n^{1-\epsilon}$  for any  $\epsilon > 0$ , fixed in advance, is already an NP-hard problem [Hå99], [Zu07]. Moreover, modulo some plausible complexity assumptions,

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it is hard to approximate the highest density of an  $m$ -subset for a given  $m$ , within a constant factor, fixed in advance [Bh12]. The best known efficient approximation achieves the factor of  $n^{1/4}$  in quasi-polynomial  $n^{O(\ln n)}$  time [B+10]. There are indications that the factor  $n^{1/4}$  might be hard to beat [B+12]. We note that the most interesting case is when  $m$  grows and  $n \gg m$ , since the highest density of an  $m$ -subset can be computed in polynomial time up to an additive error of  $\epsilon n^2/m^2$  for any  $\epsilon > 0$ , fixed in advance [FK99] (and if  $m$  is fixed in advance, the densest  $m$ -subset can be found by the exhaustive search in polynomial time).

**(1.1) Partition function.** In this paper, we approach the problem of finding the densest, or just a reasonably dense subset, via computing the *partition function*

$$(1.1.1) \quad \text{den}_m(G; \gamma) = \binom{n}{m}^{-1} \sum_{\substack{S \subset V: \\ |S|=m}} \exp \{ \gamma m \sigma(S) \},$$

where  $\gamma > 0$  is a parameter. We are interested in computing (approximating)  $\text{den}_m(G; \gamma)$  efficiently. The *exponential tilting*,  $\sigma(S) \mapsto \exp \{ \gamma m \sigma(S) \}$ , see for example, Section 13.7 of [Te99], puts greater emphasis on the sets of higher density. Let us consider the set  $\binom{V}{m}$  of all  $m$ -subsets of  $V$  as a probability space with the uniform measure. By the Markov inequality, for any  $0 < \sigma_0 < 1$ , we have

$$(1.1.2) \quad \sigma_0 + \frac{\ln \mathbf{P}(\sigma(S) \geq \sigma_0)}{\gamma m} \leq \frac{\ln \text{den}_m(G; \gamma)}{\gamma m} \leq \max_{\substack{S \subset V: \\ |S|=m}} \sigma(S),$$

so the larger  $\gamma$  we can afford, the better approximation for the densest  $m$ -subset we get. In particular, if we could choose  $\gamma \gg \ln n$  then from (1.1.2) we could approximate the highest density of an  $m$ -subset within an arbitrarily small additive error.

The partition function (1.1.1) was introduced in [Ba15], where an algorithm of quasi-polynomial  $n^{O(\ln m - \ln \epsilon)}$  complexity was constructed to compute (1.1.1) within relative error  $0 < \epsilon < 1$ , when  $\gamma = 0.07$  and when  $\gamma = 0.27$ , under additional assumptions that  $n \geq 8m$  and  $m \geq 10$ . It follows from (1.1.2) that if the probability to hit an  $m$ -subset  $S$  of density at least  $\sigma_0$  at random is  $e^{-o(m)}$  then we can certify the existence of an  $m$ -subset of density at least  $\sigma_0 - o(1)$  in quasi-polynomial time, just by computing (1.1.1). It is also shown in [Ba15] that by successive conditioning, one can find in quasi-polynomial time an  $m$ -subset  $S$  with density at least as high as certified by the value of (1.1.1).

In this paper, we present an algorithm, which, for any  $0 < \gamma < 1$ , fixed in advance, and a given  $0 < \epsilon < 1$ , computes the value of (1.1.1) within relative error  $\epsilon$  in quasi-polynomial  $n^{O(\ln m - \ln \epsilon)}$  time, provided  $n > \omega(\gamma)m$  for some constant  $\omega(\gamma) > 1$ . This improvement from  $\gamma = 0.27$  to  $\gamma = 1$  makes the algorithm competitive in some situations where it was not competitive before. Suppose, for example, we want to separate efficiently the graphs that have sufficiently many  $m$ -cliques

from the graphs that are sufficiently far from having a single  $m$ -clique. Below we show that for  $\gamma < 0.5$  our algorithm is inferior to a simple test based on the Kruskal - Katona Theorem, while for  $\gamma > 0.5$  the former can cover a greater range than the latter.

**(1.2) Example: testing graphs for  $m$ -cliques.** Let us fix two numbers  $0 < \delta < 1$  and  $\alpha > 0$  and consider the following two mutually exclusive conditions:

(1.2.1) For every  $S \subset V$  such that  $|S| = m$  we have  $\sigma(S) \leq 1 - \delta$

and

(1.2.2) If  $S \subset V$  is a random subset, sampled uniformly from the set  $\binom{V}{m}$  of all  $m$ -sets of vertices, then the probability that  $S$  is a clique is at least  $e^{-\alpha m}$ .

Suppose further, we are presented with a graph  $G = (V, E)$  and told that either condition (1.2.1) or condition (1.2.2) holds. Our goal is to decide which one. This is somewhat in the spirit of “property testing” [Go17].

We observe that if (1.2.1) holds then  $\text{den}_m(G; \gamma) \leq e^{\gamma m(1-\delta)}$  and if (1.2.2) holds then  $\text{den}_m(G; \gamma) \geq e^{(\gamma-\alpha)m}$ . Consequently, if

$$(1.2.3) \quad \alpha < \gamma\delta$$

and we can approximate  $\text{den}_m(G; \gamma)$  efficiently, we can efficiently tell (1.2.1) and (1.2.2) apart.

An anonymous referee to [Ba15] noticed that another, much simpler, algorithm can be inferred from the Kruskal - Katona Theorem. Let  $|V| = n$ . If (1.2.1) holds then  $|E| \leq (1-\delta)\binom{n}{2}$ . The Kruskal - Katona Theorem (see, for example, Section 5 of [Bo86]) implies that if (1.2.2) holds, then for every  $k$  such that  $\binom{k}{m} \leq e^{-\alpha m} \binom{n}{m}$ , we must have  $|E| \geq \binom{k}{2}$ , the model case being a graph  $G$  consisting of a  $k$ -clique and  $n - k$  isolated vertices. A computation shows that as  $n \rightarrow \infty$ , we can tell (1.2.1) and (1.2.2) apart just by counting the edges of  $G$ , provided

$$(1.2.4) \quad \alpha < -\frac{1}{2} \ln(1 - \delta).$$

Comparing (1.2.3) and (1.2.4), we observe that the algorithm based on computing the partition function  $\text{den}_m(G; \gamma)$  is not competitive as long as  $\gamma < 0.5$ , which is the case in [Ba15], but becomes competitive at least for small values of  $\delta$  as soon as  $\gamma > 0.5$ . Numerical estimates show that as long as we can choose  $\gamma > 0$  arbitrarily close 1, the condition (1.2.3) serves a wider range of  $\alpha$  than the condition (1.2.4) provided  $\delta < 0.7968$ .

We still don’t know, however, if (1.1.1) can be efficiently computed for *any*  $\gamma > 0$ , fixed in advance, and as we remarked above, it is unlikely that (1.1.1) can be efficiently computed for  $\gamma \gg \ln n$ . Our numerical experiments seem to indicate that we can afford a substantially larger  $\gamma$ . This can be partially explained by the

fact that for the Erdős - Rényi random graph  $G(n, 0.5)$  indeed a much larger  $\gamma$  can be used with high probability, see Theorem 1.5 below.

The improvement from  $\gamma = 0.27$  to an arbitrary  $\gamma < 1$  required the addition of some new ideas to the technique of [Ba15]. The approach of [Ba15] and of this paper are based on the “interpolation method” [Ba16]. As applied to our case, the idea of the method is to consider  $\text{den}_m(G; z)$  for a *complex* parameter  $z$ . We can efficiently approximate  $\text{den}_m(G; z)$  at  $z = \gamma$  if there is a connected open set  $U \subset \mathbb{C}$ , not dependent on  $m$  or  $G$ , such that  $0 \in U$ ,  $\gamma \in U$  and  $\text{den}_m(G; z) \neq 0$  for all  $z \in U$ . In [Ba15], the set  $U$  is a disk centered at  $z = 0$ , whereas in the current paper it is a thin neighborhood of the interval  $[0, \gamma]$ , which allows us to reach larger  $\gamma$ , but also requires a more refined analysis to establish zero-freeness. We give some more details now.

**(1.3) Multivariate partition function.** Given  $n \times n$  symmetric complex matrix  $Z = (z_{ij})$  and  $2 \leq m \leq n$ , we define

$$(1.3.1) \quad P_m(Z) = \sum_{\substack{S \subset \{1, \dots, n\} \\ |S|=m}} \exp \left\{ \sum_{\substack{\{i, j\} \subset S \\ i \neq j}} z_{ij} \right\}.$$

Note that the diagonal entries of  $Z$  are irrelevant, so we assume that  $z_{ii} = 0$  for all  $i$ .

Given a graph  $G = (V, E)$  with set  $V = \{1, \dots, n\}$  of vertices and  $\gamma > 0$ , we define  $Z_0 = (z_{ij})$  by

$$z_{ij} = \begin{cases} \frac{\gamma}{m-1} & \text{if } \{i, j\} \in E \\ -\frac{\gamma}{m-1} & \text{if } \{i, j\} \notin E, \end{cases}$$

and observe that

$$(1.3.2) \quad \begin{aligned} P_m(Z_0) &= \sum_{\substack{S \subset \{1, \dots, n\} \\ |S|=m}} \exp \left\{ m\gamma\sigma(S) - \frac{\gamma m}{2} \right\} \\ &= \exp \left\{ -\frac{\gamma m}{2} \right\} \binom{n}{m} \text{den}_m(G; \gamma). \end{aligned}$$

Hence to compute (1.1.1) it suffices to compute  $P_m(Z_0)$ . We compute  $P_m(Z_0)$  by interpolation, see [Ba15], [Ba16]. For that, it suffices to show that  $P_m(Z) \neq 0$  in some neighborhood of a path connecting the zero matrix to  $Z_0$  in the space of complex matrices.

We prove the following result.

**(1.4) Theorem.** *For any  $0 < \delta < 1$  there exist  $\eta = \eta(\delta) > 0$  and  $\omega = \omega(\delta) > 1$  such that if  $n \geq \omega m$  then  $P_m(Z) \neq 0$  for any  $n \times n$  symmetric complex matrix  $Z = (z_{ij})$  such that*

$$|\Re z_{ij}| \leq \frac{\delta}{m-1} \quad \text{and} \quad |\Im z_{ij}| \leq \frac{\eta}{m-1} \quad \text{for all } 1 \leq i \neq j \leq n.$$

We prove Theorem 1.4 in Sections 2 and 3. Using Theorem 1.4, in Section 4 we present an algorithm of quasi-polynomial  $n^{O(\ln m)}$  complexity to compute  $P_m(Z_0)$  and hence  $\text{den}_m(G; \gamma)$  for any  $0 < \gamma < 1$ , fixed in advance.

In [Ba15] it was established that  $P_m(Z) \neq 0$  in a polydisc

$$\mathcal{D}_{m,n} = \left\{ Z = (z_{ij}) : |z_{ij}| \leq \frac{0.27}{m-1} \text{ for all } 1 \leq i \neq j \leq n \right\}$$

provided  $n \gg m$  and  $m$  is large enough. In Theorem 1.4, we establish that  $P_m(Z) \neq 0$  in a more “economical” domain, “stretched” along the real part of the complex space of matrices. This allows us to improve the constant  $\gamma$  for which  $\text{den}_m(G; \gamma)$  is still efficiently computable.

In Section 5, we discuss some results of our numerical experiments, which seem to indicate that we can afford an essentially bigger  $\delta$  in Theorem 1.4. This can be partially explained by the fact that for the Erdős - Rényi random graph  $G(n, 0.5)$  this is indeed the case. Namely, we prove the following result in Section 6.

**(1.5) Theorem.** *Let us choose positive integers  $n$  and  $2 \leq m \leq n$ . For  $n \times n$  symmetric matrix  $W = (w_{ij})$  of independent random variables, where*

$$\mathbf{P}(w_{ij} = 1) = \mathbf{P}(w_{ij} = -1) = \frac{1}{2},$$

*we define the polynomial*

$$h_W(z) = \binom{n}{m}^{-1} \sum_{\substack{S \subset \{1, \dots, n\} \\ |S|=m}} \prod_{\{i,j\} \subset S} (1 + zw_{ij}).$$

*Let  $r > 0$  and  $\tau > 1$  be real numbers. If  $n \geq 2m^2(1+r^2)^m + 2m$  then the probability that  $h_W(z)$  has a root in the disc  $|z| < r/\sqrt{2\tau}$  does not exceed  $1/\tau$ .*

In particular, if  $n \gg m^2$  then with high probability  $h_W(z)$  has no roots in the disc  $|z| < c/\sqrt{m}$ , for an arbitrary large  $c > 0$ , fixed in advance. Similarly, if  $\ln n \gg m$  then with high probability  $h_W(z)$  has no roots in the disc  $|z| < c$  for an arbitrary large  $c > 0$ , fixed in advance.

The polynomial  $h_W(z)$  is easily translated into the partition function  $\text{den}_m(G; \gamma)$ , where  $G$  is the graph with set  $V = \{1, \dots, n\}$  of vertices and two vertices  $\{i, j\}$  span an edge if and only if  $w_{ij} = 1$ : for  $0 < \alpha < 1$ , we have

$$(1.5.1) \quad h_W(\alpha) = (1 - \alpha)^{\binom{m}{2}} \text{den}_m(G; \gamma) \quad \text{where} \quad \gamma = \frac{m-1}{2} \ln \frac{1+\alpha}{1-\alpha}.$$

Consequently, with high probability we can approximate  $\text{den}_m(G; \gamma)$  in quasi-polynomial time for  $\gamma$  as large as  $\gamma = \sqrt{m}$  provided  $n \gg m^2$  and as large as  $\gamma = m$

provided  $\ln n \gg m$ . Since the graphs we experimented on were to a large degree random (but not necessarily Erdős - Rényi  $G(n, 0.5)$ ), we may have obtained overly optimistic numerical evidence.

As is easily seen,  $\mathbf{E} h_W(\alpha) = 1$  and from our proof in Section 6 it follows that  $h_W(\alpha)$  is strongly concentrated. For example, in the regime of  $n = \Omega(m^2)$  and  $\alpha = 1/\sqrt{m}$ , we have  $\mathbf{var} h_W(\alpha) = O(1)$ . This concentration, however, does not allow us to predict with high probability the value of  $h_W(\alpha)$  with the precision that the interpolation technique based on Theorem 1.5 allows for.

In Section 6, we also discuss what may happen if  $G$  is a random graph  $G(n, 0.5)$  with a planted  $m$ -clique.

## 2. PRELIMINARIES

We consider the partition function  $P_m$  of Section 1.3 within a family of partition functions, which will allow us to prove Theorem 1.4 by induction.

**(2.1) Functionals  $P_\Omega(Z)$ .** Let us fix integers  $n$  and  $2 \leq m \leq n$ . For a subset  $\Omega \subset \{1, \dots, n\}$  and  $n \times n$  complex symmetric matrix  $Z = (z_{ij})$ , we define

$$P_\Omega(Z) = \sum_{\substack{S \subset \{1, \dots, n\}: \\ |S|=m, \Omega \subset S}} \exp \left\{ \sum_{\substack{\{i,j\} \subset S \\ i \neq j}} z_{ij} \right\}$$

where we agree that  $P_\Omega(Z) = 0$  if  $|\Omega| > m$ . In other words, we restrict the sum (1.3.1) defining  $P_m(Z)$  onto subsets  $S$  containing a given set  $\Omega$ . In particular,

$$P_\Omega(Z) = P_m(Z) \quad \text{if } \Omega = \emptyset.$$

The induction will be built on the following straightforward formulas:

$$(2.1.1) \quad P_\Omega(Z) = \frac{1}{m - |\Omega|} \sum_{j \in \{1, \dots, n\} \setminus \Omega} P_{\Omega \cup \{j\}}(Z) \quad \text{provided } |\Omega| < m$$

and for  $i \neq j$ , we have

$$(2.1.2) \quad \frac{\partial}{\partial z_{ij}} P_\Omega(Z) = \begin{cases} P_\Omega(Z) & \text{if } i, j \in \Omega, \\ P_{\Omega \cup \{j\}}(Z) & \text{if } i \in \Omega, j \notin \Omega, \\ P_{\Omega \cup \{i\}}(Z) & \text{if } i \notin \Omega, j \in \Omega, \\ P_{\Omega \cup \{i, j\}}(Z) & \text{if } i, j \notin \Omega. \end{cases}$$

We will often consider complex numbers as vectors in the plane, by identifying  $\mathbb{C} = \mathbb{R}^2$  and measuring, in particular, angles between non-zero complex numbers. We will use the following geometric lemma.

**(2.2) Lemma.** Let  $u_1, \dots, u_n \in \mathbb{C}$  be non-zero complex numbers such that the angle between any two does not exceed  $\theta$  for some  $0 < \theta < \pi/2$ . Suppose that

$$\Im \left( \sum_{j=1}^n u_j \right) = 0 \quad \text{and} \quad \sum_{j=1}^n |u_j| = c.$$

Then

$$\sum_{j=1}^n |\Im u_j| \leq c \sin \frac{\theta}{2}.$$

*Proof.* Scaling  $u_j$ , if necessary, without loss of generality we assume that  $c = 1$ .

Without loss of generality, we assume that  $\arg u_j \neq 0$  for  $j = 1, \dots, n$ . Indeed, if  $\arg u_j = 0$  for some  $j$ , we can remove the vector from the collection, which would make the sum

$$(2.2.1) \quad \sum_{j=1}^n |u_j|$$

only smaller. Rescaling  $u_j \mapsto \tau u_j$  for some real  $\tau > 1$ , we make (2.2.1) equal to 1 and increase

$$(2.2.2) \quad \sum_{j=1}^n |\Im u_j|.$$

Reflecting the vectors  $u_j$  in the coordinate axes if necessary, without loss of generality we may assume that  $\Re u_1 \geq 0$  and  $\Im u_1 > 0$ . Hence there is a vector, say  $u_2$ , such that  $\Im u_2 < 0$ . We necessarily have  $\Re u_2 \geq 0$ , since otherwise the angle between  $u_1$  and  $u_2$  exceeds  $\pi/2$ . Then for any vector  $u_j$ , we must have  $\Re u_j \geq 0$ , since otherwise one of the angles formed by  $u_j$  with  $u_1$  or  $u_2$  will exceed  $\pi/2$ .

Hence without loss of generality, we assume that  $\Re u_j > 0$  for  $j = 1, \dots, n$ . Let

$$\alpha = \max_{j=1, \dots, n} \arg u_j,$$

so that

$$0 < \alpha < \theta$$

and let

$$-\beta = \min_{j=1, \dots, n} \arg u_j < 0.$$

Then  $\alpha + \beta \leq \theta$ .

Let

$$J_+ = \{j : \arg u_j > 0\} \quad \text{and} \quad J_- = \{j : \arg u_j < 0\}.$$

Next, without loss of generality, we assume that  $\arg u_j = \alpha$  for all  $j \in J_+$  and that  $\arg u_j = -\beta$  for all  $j \in J_-$ . Indeed, suppose that  $\arg u_1 = \alpha_1$  where  $0 < \alpha_1 < \alpha$ . We can modify

$$u_1 \mapsto \frac{\sin \alpha_1}{\sin \alpha} e^{i(\alpha - \alpha_1)} u_1$$

(we rotate and shrink  $u_1$  so as to make its argument equal to  $\alpha$  and leave  $\Im u_1$  intact). The sum (2.2.1) gets smaller while all other conditions and the sum (2.2.2) remain intact. Rescaling  $u_j \mapsto \tau u_j$  for some real  $\tau > 1$ , we make (2.2.1) equal to 1 and increase (2.2.2), while keeping other constraints of the lemma intact. The case of  $\arg u_j > -\beta$  for some  $j \in J_-$  is handled similarly.

Next, without loss of generality, we assume that  $\alpha + \beta = \theta$ . Indeed, if  $\alpha + \beta < \theta$ , we can rotate and scale vectors  $u_j$  as above, so that the sum (2.2.2) increases while all other conditions are satisfied.

Now, let

$$u_+ = \sum_{j \in J_+} u_j \quad \text{and} \quad u_- = \sum_{j \in J_-} u_j.$$

Then  $\arg u_+ = \alpha$ ,  $\arg u_- = -\beta$ ,  $\Im(u_+ + u_-) = 0$ ,  $|u_+| + |u_-| = 1$  and (2.2.2) is equal to  $|\Im u_+| + |\Im u_-|$ .

Denoting  $a = |u_+|$  and  $b = |u_-|$ , we have  $a + b = 1$  and  $a \sin \alpha - b \sin \beta = 0$ , from which

$$a = \frac{\sin \beta}{\sin \alpha + \sin \beta} \quad \text{and} \quad b = \frac{\sin \alpha}{\sin \alpha + \sin \beta}$$

and so

$$|\Im u_+| + |\Im u_-| = \frac{2 \sin \alpha \sin \beta}{\sin \alpha + \sin \beta}.$$

Now, the function

$$\alpha \mapsto \frac{1}{\sin \alpha} \quad \text{for} \quad 0 \leq \alpha \leq \frac{\pi}{2}$$

is convex and hence the minimum of

$$\frac{\sin \alpha + \sin \beta}{\sin \alpha \sin \beta} = \frac{1}{\sin \alpha} + \frac{1}{\sin \beta}$$

on the interval  $\alpha + \beta = \theta$ ,  $\alpha, \beta \geq 0$ , is attained at  $\alpha = \beta = \theta/2$ . The proof now follows.  $\square$

We need another geometric lemma.

**(2.3) Lemma.** *Let  $u_1, \dots, u_n \in \mathbb{C}$  be non-zero complex numbers such that the angle between any two does not exceed  $\theta$  for some  $0 \leq \theta < 2\pi/3$ . Let  $u = u_1 + \dots + u_n$ . Then*

$$|u| \geq \left( \cos \frac{\theta}{2} \right) \sum_{k=1}^n |u_k|.$$

*Proof.* This is Lemma 3.1 of [Ba15] and Lemma 3.6.3 of [Ba16].  $\square$



### 3. PROOF OF THEOREM 1.4

We identify the space of  $n \times n$  zero-diagonal complex symmetric matrices  $Z = (z_{ij})$  with  $\mathbb{C}^{\binom{n}{2}}$ . Given  $\delta \geq \eta > 0$ , we define a domain  $\mathcal{U}(\delta, \eta) = \mathcal{U}_{n,m}(\delta, \eta) \subset \mathbb{C}^{\binom{n}{2}}$  by

$$\mathcal{U}(\delta, \eta) = \left\{ Z = (z_{ij}) : |\Re z_{ij}| \leq \frac{\delta}{m-1} \quad \text{and} \quad |\Im z_{ij}| \leq \frac{\eta}{m-1} \right\}.$$

If  $Z' = (z'_{ij})$  and  $Z'' = (z''_{ij})$  are two matrices from  $\mathcal{U}(\delta, \tau)$  then

$$|z'_{ij} - z''_{ij}| \leq \frac{\sqrt{(2\delta)^2 + (2\eta)^2}}{m-1} \leq \frac{2\sqrt{2}\delta}{m-1} \quad \text{for all } i, j.$$

We will prove by descending induction on  $|\Omega|$  that  $P_\Omega(Z) \neq 0$  for all  $Z \in \mathcal{U}(\delta, \eta)$  and that, moreover, a number of stronger conditions are met. The induction is based on the following two lemmas that describe how  $P_\Omega(Z)$  changes when only the entries in the  $i$ -th row and column of  $Z$  change. The first lemma deals with the case of  $i \in \Omega$ .

**(3.1) Lemma.** *Let us fix  $\Omega \subset \{1, \dots, n\}$  such that  $|\Omega| < m$ . Suppose that for any  $Z \in \mathcal{U}(\delta, \eta)$  and any  $j, k \notin \Omega$ , we have  $P_{\Omega \cup \{j\}}(Z) \neq 0$ ,  $P_{\Omega \cup \{k\}}(Z) \neq 0$  and the angle between the two non-zero complex numbers does not exceed  $\theta$  for some  $0 < \theta \leq \pi/2$ . Then*

(1) *We have*

$$P_\Omega(Z) \neq 0 \quad \text{for all } Z \in \mathcal{U}(\delta, \eta).$$

(2) *Suppose additionally, that  $\Omega \neq \emptyset$  and let us fix an  $i \in \Omega$ . Let  $Z', Z'' \in \mathcal{U}(\delta, \eta)$  be two matrices that differ only in the coordinates  $z_{ij} = z_{ji}$  for  $j \neq i$ . Then*

$$\left| \frac{P_\Omega(Z')}{P_\Omega(Z'')} \right| \leq e^{6\delta}$$

*and the angle between  $P_\Omega(Z') \neq 0$  and  $P_\Omega(Z'') \neq 0$  does not exceed*

$$2\delta \tan \frac{\theta}{2} + 5\eta.$$

*Proof.* It follows from (2.1.1) and Lemma 2.3 that

$$(3.1.1) \quad |P_\Omega(Z)| \geq \frac{\cos(\theta/2)}{m-|\Omega|} \sum_{j \notin \Omega} |P_{\Omega \cup \{j\}}(Z)| \geq \frac{1}{(m-1)\sqrt{2}} \sum_{j \notin \Omega} |P_{\Omega \cup \{j\}}(Z)|.$$

In particular, Part (1) follows.

To prove Part (2), let us choose a branch of  $\ln P_\Omega(Z)$  for  $Z \in \mathcal{U}(\delta, \eta)$ . For  $0 \leq t \leq 1$ , let  $Z(t) = tZ'' + (1-t)Z'$ . Then

$$\begin{aligned} \ln P_\Omega(Z'') - \ln P_\Omega(Z') &= \int_0^1 \frac{d}{dt} \ln P_\Omega(Z(t)) dt \\ &= \int_0^1 \sum_{j: j \neq i} (z''_{ij} - z'_{ij}) \frac{\partial}{\partial z_{ij}} \ln P_\Omega(Z) \Big|_{Z=Z(t)} dt. \end{aligned}$$

Using (2.1.2), we conclude that

$$\frac{\partial}{\partial z_{ij}} \ln P_\Omega(Z) = \begin{cases} 1 & \text{if } j \in \Omega, \\ P_{\Omega \cup \{j\}}(Z)/P_\Omega(Z) & \text{if } j \notin \Omega, \end{cases}$$

and hence

$$(3.1.2) \quad \begin{aligned} \ln P_\Omega(Z'') - \ln P_\Omega(Z') &= \sum_{j \in \Omega, j \neq i} (z''_{ij} - z'_{ij}) \\ &\quad + \int_0^1 \sum_{j \notin \Omega} (z''_{ij} - z'_{ij}) \frac{P_{\Omega \cup \{j\}}(Z(t))}{P_\Omega(Z(t))} dt. \end{aligned}$$

Using (3.1.1), we get from (3.1.2) that

$$\begin{aligned} |\Re \ln P_\Omega(Z'') - \Re \ln P_\Omega(Z')| &\leq 2\delta + (m-1)\sqrt{2} \max_{j \notin \Omega} |z''_{ij} - z'_{ij}| \\ &\leq 2\delta + 4\delta = 6\delta \end{aligned}$$

and hence

$$\left| \frac{P_\Omega(Z')}{P_\Omega(Z'')} \right| \leq e^{6\delta},$$

as claimed.

From (2.1.1), for all  $Z \in \mathcal{U}(\delta, \eta)$  we have that

$$\sum_{j \notin \Omega} \frac{P_{\Omega \cup \{j\}}(Z)}{P_\Omega(Z)} = m - |\Omega|$$

is real, while from (3.1.1), we conclude that

$$\sum_{j \notin \Omega} \left| \frac{P_{\Omega \cup \{j\}}(Z)}{P_\Omega(Z)} \right| \leq \frac{m - |\Omega|}{\cos(\theta/2)} \leq \frac{m-1}{\cos(\theta/2)}.$$

Applying Lemma 2.2 with  $u_j = P_{\Omega \cup \{j\}}(Z)/P_\Omega(Z)$ , we conclude that

$$\sum_{j \notin \Omega} \left| \Im \frac{P_{\Omega \cup \{j\}}(Z)}{P_\Omega(Z)} \right| \leq (m-1) \tan \frac{\theta}{2}.$$

Therefore, from (3.1.2),

$$\begin{aligned}
|\Im \ln P_\Omega(Z'') - \Im \ln P_\Omega(Z')| &\leq 2\eta + (m-1) \tan \frac{\theta}{2} \max_{j \notin \Omega} |\Re z''_{ij} - \Re z'_{ij}| \\
&\quad + (m-1) \sqrt{2} \max_{j \notin \Omega} |\Im z''_{ij} - \Im z'_{ij}| \\
&\leq 2\delta \tan \frac{\theta}{2} + 5\eta.
\end{aligned}$$

Hence the angle between  $P_\Omega(Z'')$  and  $P_\Omega(Z')$  does not exceed  $2\delta \tan \frac{\theta}{2} + 5\eta$ , as claimed.  $\square$

The second lemma shows that  $P_\Omega(Z)$  does not change much if only the entries of  $Z$  in the  $i$ -th row and column are changed for some  $i \notin \Omega$ , assuming that  $n \gg m$ .

**(3.2) Lemma.** *Let us fix an  $\Omega \subset \{1, \dots, n\}$ ,  $|\Omega| \leq m-1$ . Suppose for any  $i, j \notin \Omega$  and all  $Z \in \mathcal{U}(\delta, \eta)$  we have  $P_{\Omega \cup \{i\}}(Z) \neq 0$ ,  $P_{\Omega \cup \{j\}}(Z) \neq 0$  and the angle between the two complex numbers does not exceed  $\pi/2$  and that*

$$\left| \frac{P_{\Omega \cup \{i\}}(Z)}{P_{\Omega \cup \{j\}}(Z)} \right| \leq \lambda$$

for some  $\lambda \geq 1$ .

*In addition, suppose that if  $|\Omega| \leq m-2$  then for any distinct  $i, j, k \notin \Omega$  and all  $Z \in \mathcal{U}(\delta, \eta)$  we have  $P_{\Omega \cup \{i, j\}}(Z) \neq 0$ ,  $P_{\Omega \cup \{i, k\}}(Z) \neq 0$  and the angle between the two complex numbers does not exceed  $\pi/2$ .*

*Let us fix an  $i \notin \Omega$  and let  $Z', Z'' \in \mathcal{U}(\delta, \eta)$  be two matrices that differ only in the coordinates  $z_{ij} = z_{ji}$  for  $j \neq i$ . Then*

$$\left| \frac{P_\Omega(Z')}{P_\Omega(Z'')} \right| \leq \exp \left\{ \frac{10\delta\lambda m}{n-1} \right\}$$

*and the angle between  $P_\Omega(Z') \neq 0$  and  $P_\Omega(Z'') \neq 0$  does not exceed*

$$\frac{10\delta\lambda m}{n-1}.$$

*Proof.* It follows from Lemma 3.1 that  $P_\Omega(Z) \neq 0$  for all  $Z \in \mathcal{U}(\delta, \eta)$ .

Arguing as in the proof of Lemma 3.1, we introduce  $Z(t) = tZ'' + (1-t)Z'$  and write

$$\ln P_\Omega(Z'') - \ln P_\Omega(Z') = \int_0^1 \sum_{j: j \neq i} (z''_{ij} - z'_{ij}) \frac{\partial}{\partial z_{ij}} \ln P_\Omega(Z) \Big|_{Z=Z(t)} dt.$$

From (2.1.2), we write

$$(3.2.1) \quad \begin{aligned} \ln P_\Omega(Z'') - \ln P_\Omega(Z') &= \int_0^1 \sum_{j \in \Omega} (z''_{ij} - z'_{ij}) \frac{P_{\Omega \cup \{i\}}(Z(t))}{P_\Omega(Z(t))} \\ &+ \sum_{j \notin \Omega, j \neq i} (z''_{ij} - z'_{ij}) \frac{P_{\Omega \cup \{i, j\}}(Z(t))}{P_\Omega(Z(t))} dt. \end{aligned}$$

Suppose first that  $|\Omega| \leq m - 2$ . From (2.1.1), we have

$$P_{\Omega \cup \{i\}}(Z) = \frac{1}{m - |\Omega| - 1} \sum_{j \notin \Omega, j \neq i} P_{\Omega \cup \{i, j\}}(Z).$$

Applying Lemma 2.3, we get that

$$(3.2.2) \quad \sum_{j \notin \Omega, j \neq i} |P_{\Omega \cup \{i, j\}}(Z)| \leq (m - 1)\sqrt{2} |P_{\Omega \cup \{i\}}(Z)|$$

for all  $Z \in \mathcal{U}(\delta, \eta)$ .

Since by (2.1.1) we also have

$$P_\Omega(Z) = \frac{1}{m - |\Omega|} \sum_{j \notin \Omega} P_{\Omega \cup \{j\}}(Z),$$

applying Lemma 2.3, we conclude that

$$\sum_{j \notin \Omega} |P_{\Omega \cup \{j\}}(Z)| \leq (m - |\Omega|)\sqrt{2} |P_\Omega(Z)|.$$

Hence for all  $i \notin \Omega$ , we have

$$(3.2.3) \quad |P_{\Omega \cup \{i\}}(Z)| \leq \frac{\lambda(m - |\Omega|)\sqrt{2}}{n - |\Omega|} |P_\Omega(Z)| \leq \frac{\lambda m \sqrt{2}}{n} |P_\Omega(Z)|.$$

Combining (3.2.3) and (3.2.2), we get

$$(3.2.4) \quad \sum_{j \notin \Omega, j \neq i} |P_{\Omega \cup \{i, j\}}(Z)| \leq \frac{2\lambda m(m - 1)}{n} |P_\Omega(Z)|.$$

Combining (3.2.1), (3.2.2), (3.2.3) and (3.2.4), we get

$$\begin{aligned} |\ln P_\Omega(Z'') - \ln P_\Omega(Z')| &\leq \frac{2\sqrt{2}\delta}{m - 1} \cdot \frac{\lambda|\Omega|(m - |\Omega|)\sqrt{2}}{n - |\Omega|} + \frac{2\sqrt{2}\delta}{m - 1} \cdot \frac{2\lambda m(m - 1)}{n} \\ &\leq \frac{4\delta\lambda m}{n - 1} + \frac{4\sqrt{2}\delta\lambda m}{n} \leq \frac{10\delta\lambda m}{n - 1}. \end{aligned}$$

If  $|\Omega| = m - 1$  then from (3.2.1) and (3.2.3), we get

$$|\ln P_\Omega(Z'') - \ln P_\Omega(Z')| \leq \frac{2\sqrt{2}\delta}{m - 1} \cdot \frac{\lambda m \sqrt{2}}{n} \leq \frac{4\delta\lambda m}{n - 1},$$

which concludes the proof.  $\square$

Now we are ready to prove Theorem 1.4.

**(3.3) Proof of Theorem 1.4.** Given  $0 < \delta < 1$ , we choose  $0 < \theta < \pi/2$  so that

$$2\delta \tan \frac{\theta}{2} < \theta.$$

We then choose  $\eta > 0$  such that

$$2\delta \tan \frac{\theta}{2} + 5\eta < \theta.$$

We choose

$$\lambda > e^{6\delta}$$

and choose  $\omega > 1$  so that

$$2\delta \tan \frac{\theta}{2} + 5\eta + \frac{10\delta\lambda m}{n-1} \leq \theta \quad \text{and} \quad \exp \left\{ 6\delta + \frac{10\delta\lambda m}{n-1} \right\} \leq \lambda$$

whenever  $n \geq \omega m$ .

Suppose that  $n \geq \omega m$ . We prove by descending induction on  $r = m, m-1, \dots, 1$  that if  $\Omega_1, \Omega_2 \in \{1, \dots, n\}$  are two sets such that  $|\Omega_1| = |\Omega_2| = r$  and  $|\Omega_1 \Delta \Omega_2| = 2$  then for all  $Z \in \mathcal{U}(\delta, \eta)$  we have  $P_{\Omega_1}(Z) \neq 0$ ,  $P_{\Omega_2}(Z) \neq 0$ , the angle between  $P_{\Omega_1}(Z)$  and  $P_{\Omega_2}(Z)$  does not exceed  $\theta$  while the ratio of  $|P_{\Omega_1}(Z)|$  and  $|P_{\Omega_2}(Z)|$  does not exceed  $\lambda$ .

Assume that  $r = m$ . Without loss of generality, we assume that  $\Omega_1 = \Omega \cup \{1\}$  and  $\Omega_2 = \Omega \cup \{2\}$  for some  $\Omega \subset \{3, \dots, n\}$  such that  $|\Omega| = m-1$ . We have

$$P_{\Omega_1}(Z) = \exp \left\{ \sum_{\{i,j\} \subset \Omega} z_{ij} \right\} \exp \left\{ \sum_{i \in \Omega} z_{1i} \right\} \quad \text{and}$$

$$P_{\Omega_2}(Z) = \exp \left\{ \sum_{\{i,j\} \subset \Omega} z_{ij} \right\} \exp \left\{ \sum_{i \in \Omega} z_{2i} \right\}.$$

Clearly,  $P_{\Omega_1}(Z) \neq 0$ ,  $P_{\Omega_2}(Z) \neq 0$ , the angle between  $P_{\Omega_1}(Z)$  and  $P_{\Omega_2}(Z)$  does not exceed  $2\eta \leq \theta$  while the ratio of  $|P_{\Omega_1}(Z)|$  and  $|P_{\Omega_2}(Z)|$  does not exceed  $e^{2\delta} \leq \lambda$ .

Suppose now that the statements hold for all subsets  $\Omega \subset \{1, \dots, n\}$  of cardinality at least  $r+1$  for some  $r \leq m-1$  and let  $\Omega_1, \Omega_2 \subset \{1, \dots, n\}$  be two subsets of cardinality  $r \geq 1$  such that  $|\Omega_1 \Delta \Omega_2| = 2$ . Again, without loss of generality, we assume that  $\Omega_1 = \Omega \cup \{1\}$  and  $\Omega_2 = \Omega \cup \{2\}$  for some  $\Omega \subset \{3, \dots, n\}$  such that  $|\Omega| = r-1$ . Then we observe that  $P_{\Omega_2}(Z) = P_{\Omega_1}(Z')$ , where

$$z'_{1i} = z'_{i1} = z_{2i} = z_{i2} \quad \text{and} \quad z'_{2i} = z'_{i2} = z_{1i} = z_{i1} \quad \text{for } i \neq 1, 2,$$

while all other entries of  $Z$  and  $Z'$  coincide. Applying Lemma 3.1 and Lemma 3.2 and the induction hypothesis to sets  $\Omega_1 \cup \{j\}$  for  $j \notin \Omega_1$  and  $\Omega_1 \cup \{j, k\}$  for

$j, k \notin \Omega_1$ , we conclude that the angle between  $P_{\Omega_1}(Z) \neq 0$  and  $P_{\Omega_2}(Z) \neq 0$  does not exceed

$$2\delta \tan \frac{\theta}{2} + 5\eta + \frac{10\delta\lambda m}{n-1} \leq \theta,$$

while the ratio of  $|P_{\Omega_1}(Z)|$  and  $|P_{\Omega_2}(Z)|$  does not exceed

$$\exp \left\{ 6\delta + \frac{10\delta\lambda m}{n-1} \right\} \leq \lambda.$$

This proves that  $P_{\{i\}}(Z) \neq 0$  for all  $i \in \{1, \dots, n\}$  and all  $Z \in \mathcal{U}(\delta, \eta)$  and that the angle between  $P_{\{i\}}(Z) \neq 0$  and  $P_{\{j\}}(Z) \neq 0$  does not exceed  $\theta$  for all  $i, j \in \{1, \dots, n\}$ . From (2.1.1) we conclude that  $P_m(Z) = P_\emptyset(Z) \neq 0$  for all  $Z \in \mathcal{U}(\delta, \eta)$ .  $\square$

#### 4. COMPUTING THE PARTITION FUNCTION

Here we show how to compute the density partition function  $\text{den}_m(G; \gamma)$ . First, we make a change of coordinates to convert the partition function  $P_m(Z)$  of Section 1.3 into a multivariate polynomial.

**(4.1) A polynomial version of  $P_m(Z)$ .** For an  $n \times n$  complex symmetric matrix  $W = (w_{ij})$  with zero diagonal, we define

$$p_m(W) = \binom{n}{m}^{-1} \sum_{\substack{S \subseteq \{1, \dots, n\} \\ |S|=m}} \prod_{\substack{\{i, j\} \subseteq S \\ i \neq j}} (1 + w_{ij}).$$

Hence  $p_m(W)$  is a polynomial of degree  $\binom{m}{2}$  in the entries  $w_{ij}$  and, assuming that  $|w_{ij}| < 1$  for all  $i, j$ , we can write

$$p_m(W) = \binom{n}{m}^{-1} P_m(Z) \quad \text{where} \quad Z = (z_{ij}) \quad \text{and} \quad z_{ij} = \ln(1 + w_{ij})$$

(we choose the standard branch of the logarithm in the right half-plane of  $\mathbb{C}$ ). Theorem 1.4 implies that for every  $0 < \delta < 1$  there is  $\eta = \eta(\delta) > 0$  and  $\omega = \omega(\delta) > 1$  such that

$$(4.1.1) \quad \begin{aligned} p_m(W) \neq 0 \quad \text{whenever} \quad & |\Re \ln(1 + w_{ij})| \leq \frac{\delta}{m-1}, \\ & |\Im \ln(1 + w_{ij})| \leq \frac{\eta}{m-1} \quad \text{and} \\ & n \geq \omega m. \end{aligned}$$

To compute  $\text{den}_m(G; \gamma)$  for a given  $0 < \gamma < 1$  and a given graph  $G = (V, E)$ , we define

$$(4.1.2) \quad w_{ij} = \begin{cases} \exp \left\{ \frac{\gamma}{m-1} \right\} - 1 & \text{if } \{i, j\} \in E, \\ \exp \left\{ -\frac{\gamma}{m-1} \right\} - 1 & \text{if } \{i, j\} \notin E. \end{cases}$$

Then, by (1.3.2), we have

$$(4.1.3) \quad \text{den}_m(G; \gamma) = \exp \left\{ \frac{\gamma m}{2} \right\} p_m(W).$$

The interpolation method is based on the following simple lemma.

**(4.2) Lemma.** *Let  $g : \mathbb{C} \rightarrow \mathbb{C}$  be a univariate polynomial and suppose that  $g(z) \neq 0$  provided  $|z| < \beta$  where  $\beta > 1$  is some real number. Let us choose a branch of  $f(z) = \ln g(z)$  in the disc  $|z| < \beta$  and let*

$$T_r(z) = f(0) + \sum_{k=1}^r \frac{f^{(k)}(0)}{k!} z^k$$

be the Taylor polynomial of  $f$  of degree  $r$  computed at  $z = 0$ . Then

$$|f(1) - T_r(1)| \leq \frac{\deg g}{\beta^r (\beta - 1)(r + 1)}.$$

*Proof.* This is Lemma 2.2.1 of [Ba16], see also Lemma 1.1 of [Ba15].  $\square$

The gist of Lemma 4.2 is that to approximate  $f(1)$  within an additive error  $\epsilon$ , it suffices to compute the Taylor polynomial of  $f(z)$  at 0 of degree  $r = O_\beta(\ln \deg g - \ln \epsilon)$ , where the implicit constant in the “ $O$ ” notation depends on  $\beta$  alone. We would like to apply Lemma 4.2 to the univariate polynomial

$$(4.2.1) \quad h(z) = \binom{n}{m}^{-1} \sum_{\substack{S \subset \{1, \dots, n\} \\ |S|=m}} \prod_{\substack{\{i, j\} \subset S \\ i \neq j}} (1 + zw_{ij}),$$

where  $w_{ij}$  are defined by (4.1.2). Indeed, the value we are ultimately interested is  $h(1) = p_m(W)$ . However, Lemma 4.2 requires that  $h(z) \neq 0$  in a disc of some radius  $\beta > 1$ , whereas (4.1.1) only guarantees that  $h(z) \neq 0$  for  $z$  in a neighborhood of the interval  $[0, 1] \subset \mathbb{C}$ . To remedy this, we compose  $h$  with a polynomial  $\phi : \mathbb{C} \rightarrow \mathbb{C}$  such that  $\phi(0) = 0$ ,  $\phi(1) = 1$  and  $\phi$  maps the disc  $|z| < \beta$  for some  $\beta > 1$  inside the prescribed neighborhood of  $[0, 1] \subset \mathbb{C}$ . We then apply Lemma 4.2 to the composition  $g(z) = h(\phi(z))$ . The following lemma provides an explicit construction of  $\phi$ .

**(4.3) Lemma.** *For  $0 < \rho < 1$ , we define*

$$\alpha = \alpha(\rho) = 1 - e^{-\frac{1}{\rho}}, \quad \beta = \beta(\rho) = \frac{1 - e^{-1 - \frac{1}{\rho}}}{1 - e^{-\frac{1}{\rho}}} > 1,$$

$$N = N(\rho) = \left\lceil \left(1 + \frac{1}{\rho}\right) e^{1 + \frac{1}{\rho}} \right\rceil, \quad \sigma = \sigma(\rho) = \sum_{k=1}^N \frac{\alpha^k}{k} \quad \text{and}$$

$$\phi(z) = \phi_\rho(z) = \frac{1}{\sigma} \sum_{k=1}^N \frac{(\alpha z)^k}{k}.$$

Then  $\phi : \mathbb{C} \rightarrow \mathbb{C}$  is a polynomial of degree  $N$  such that  $\phi(0) = 0$ ,  $\phi(1) = 1$ ,

$$-\rho \leq \Re \phi(z) \leq 1 + 2\rho \quad \text{and} \quad |\Im \phi(z)| \leq 2\rho$$

provided  $|z| \leq \beta$ .

*Proof.* This is Lemma 2.2.3 of [Ba16]. □

Lemma 4.2 also requires the derivatives  $f^{(k)}(0)$  of  $f(z) = \ln g(z)$  at  $z = 0$ . Those, however, can be easily computed from the derivatives  $g^{(k)}(0)$ , as described in Section 2.2.2 of [Ba16], see also Section 2.1 of [Ba15]. We briefly sketch how.

**(4.4) Computing derivatives.** Suppose that  $f(z) = \ln g(z)$  as in Lemma 4.2.

Then

$$f'(z) = \frac{g'(z)}{g(z)} \quad \text{and} \quad g'(z) = f'(z)g(z).$$

Differentiating the product  $k - 1$  times, we obtain

$$(4.4.1) \quad g^{(k)}(0) = \sum_{j=0}^{k-1} \binom{k-1}{j} f^{(k-j)}(0)g^{(j)}(0) \quad \text{for } k = 1, \dots, r.$$

We interpret (4.4.1) as a system of linear equations in variables  $f^{(k)}(0)$  for  $k = 1, \dots, r$  with coefficients  $g^{(k)}(0)$  for  $k = 0, \dots, r$ . This is a triangular system of linear equations with non-zero entries  $g^{(0)}(0) = g(0)$  on the diagonal, that can be solved in  $O(r^2)$  time, provided the values of  $g^{(k)}(0)$  are known.

To supply the last ingredient of the algorithm, we show how to compute  $h^{(k)}(0)$  for  $k = 0, \dots, r$ , where  $h$  is the polynomial defined by (4.2.1). This is also done in [Ba15], but we reproduce it here for completeness.

We have

$$h^{(k)}(0) = \binom{n}{m}^{-1} \sum_{\substack{S \subset \{1, \dots, n\} \\ |S|=m}} \sum_{\{i_1, j_1\}, \dots, \{i_k, j_k\} \subset S} w_{i_1 j_1} \cdots w_{i_k j_k},$$

where the inner sum is taken over all ordered collections of distinct unordered pairs  $\{i_1, j_1\}, \dots, \{i_k, j_k\} \subset S$ . For such a collection, say  $I$ , let  $\nu(I)$  be the number of distinct vertices among  $i_1, j_1, \dots, i_k, j_k$ . Then there are exactly  $\binom{n-\nu(I)}{m-\nu(I)}$  different  $m$ -subsets  $S$  containing the edges from  $I$  and we can rewrite the above sum as

$$(4.4.2) \quad h^{(k)}(0) = \binom{n}{m}^{-1} \sum_{I = (\{i_1, j_1\}, \dots, \{i_k, j_k\})} \binom{n-\nu(I)}{m-\nu(I)} w_{i_1 j_1} \cdots w_{i_k j_k},$$

where the sum is taken over all ordered collections of  $k$  unordered pairs  $\{i_s, j_s\}$ . It is clear now that  $h^{(k)}(0)$  can be computed in  $n^{O(k)}$  time by the exhaustive enumeration of all possible collections of  $k$  pairs.

In Section 5 we present faster formulas for computing  $h^{(2)}(0)$  and  $h^{(3)}(0)$  that we used for our numerical experiments.



**(4.5) The algorithm.** Let us fix  $0 < \gamma < 1$ . Below we summarize the algorithm for computing  $\text{den}_m(G; \gamma)$  within relative error  $0 < \epsilon < 1$ , by which we understand computing  $\ln \text{den}_m(G; \gamma)$  within additive error  $\epsilon$ . We assume that  $m \geq 4$  and that  $n \geq \omega m$  for some  $\omega = \omega(\gamma) > 1$ , to be specified below.

Given a graph  $G = (V, E)$  with set  $V = \{1, \dots, n\}$  of vertices, and an integer  $m \leq n$ , we compute the  $n \times n$  symmetric matrix  $W = (w_{ij})$  by (4.1.2). Since  $m \geq 4$ , we have  $|w_{ij}| \leq 0.4$  for all  $i, j$ .

Our goal is to compute  $p_m(W) = h(1)$ , where  $h$  is the univariate polynomial defined by (4.2.1). We note that  $\deg h = \binom{m}{2}$ .

Let us choose  $1 > \delta > \gamma$  and let  $\eta = \eta(\delta) > 0$  and  $\omega = \omega(\delta) > 1$  be the numbers of Theorem 1.4 and in (4.1.1). We find  $\rho = \rho(\delta) > 0$  such that

$$|\Re \ln(1 + zw_{ij})| \leq \frac{\delta}{m-1} \quad \text{and} \quad |\Im \ln(1 + zw_{ij})| \leq \frac{\eta}{m-1}$$

as long as

$$(4.5.1) \quad -\rho \leq \Re z \leq 1 + \rho \quad \text{and} \quad |\Im z| \leq \rho.$$

Indeed, if  $z \in [0, 1]$  then

$$-\frac{\gamma}{m-1} \leq \ln(1 + zw_{ij}) \leq \frac{\gamma}{m-1}$$

and for  $|z| \leq 2$ , we have

$$\left| \frac{d}{dz} \ln(1 + zw_{ij}) \right| = \left| \frac{w_{ij}}{1 + zw_{ij}} \right| \leq \frac{10}{m-1}$$

so the desired  $\rho$  can indeed be found.

It follows by (4.1.1) that  $h(z) \neq 0$  as long as  $n \geq \omega m$  and (4.5.1) holds.

Using Lemma 4.3, we construct a polynomial  $\phi : \mathbb{C} \rightarrow \mathbb{C}$  of some degree  $N = N(\rho) = N(\delta)$  such that  $\phi(0) = 0$ ,  $\phi(1) = 1$  and

$$-\rho \leq \Re \phi(z) \leq 1 + \rho \quad \text{and} \quad |\Im \phi(z)| \leq \rho$$

as long as  $|z| \leq \beta$  for some  $\beta = \beta(\rho) = \beta(\delta) > 1$ . We define

$$g(z) = h(\phi(z))$$

and our goal is to compute  $g(1) = h(\phi(1))$ . We note that

$$\deg g \leq N \deg h = N \binom{m}{2}.$$

We choose a branch of  $f(z) = \ln g(z)$  for  $z$  satisfying (4.5.1).

Using Lemma 4.2, we find an integer  $r = O_\rho(\ln m - \ln \epsilon) = O_\delta(\ln m - \ln \epsilon)$  such that

$$|T_r(1) - f(1)| \leq \epsilon,$$

where  $T_r(z)$  is the Taylor polynomial of  $f(z)$  of degree  $r$ , computed at  $z = 0$ . The implicit constant in the “ $O$ ” notation depends only on  $\rho$ , which in turn depends only on  $\delta$ . Hence our goal is to compute  $T_r(1)$ , for which we need to compute  $f^{(k)}(0)$  for  $k = 1, \dots, r$ . As in Section 4.4, we reduce it in  $O(r^2)$  time to computing  $g^{(k)}(0)$  for  $k = 1, \dots, r$ . Note that

$$g(0) = h(\phi(0)) = h(0) = 1.$$

Let  $\phi_r(z)$  be the truncation of the polynomial  $\phi(z)$  obtained by discarding all monomials of degree higher than  $r$ . Similarly, let  $h_r(z)$  be the truncation of the polynomial  $h(z)$ , obtained by discarding all monomial of degree higher than  $r$ . We compute  $h_r(z)$  as in Section 4.4 in  $n^{O(r)}$  time. Finally, we compute the truncation of the composition  $h_r(\phi_r(z))$ . A fast (polynomial in  $r$ ) way to do it, is to use Horner’s method: assuming that

$$h_r(z) = \sum_{k=0}^r b_k z^k,$$

we successively compute

$$\begin{aligned} & b_r \phi_r(z) + b_{r-1}, \quad (b_r \phi_r(z) + b_{r-1}) \phi_r(z) + b_{r-2}, \\ & ((b_r \phi_r(z) + b_{r-1}) \phi_r(z) + b_{r-2}) \phi_r(z) + b_{r-3}, \dots \end{aligned}$$

discarding on the way all monomials of degree higher than  $r$ . In the end, we have computed  $g^{(k)}(0)$  for  $k = 0, \dots, r$  and hence  $f^{(k)}(0)$  for  $k = 0, \dots, r$  and hence  $T_m(1)$  approximating  $f(1) = \ln h(1)$  within additive error  $\epsilon$ . From (4.1.3), we compute

$$\text{den}_m(G; \gamma) = \exp \left\{ \frac{\gamma m}{2} \right\} h(1)$$

within relative error  $\epsilon > 0$ .

## 5. REMARKS ON THE PRACTICAL IMPLEMENTATION

We implemented a *much* simplified version of the algorithm. Given a graph  $G = (V, E)$  with set  $V = \{1, \dots, n\}$  of vertices and an integer  $2 \leq m \leq n$ , we define the  $n \times n$  matrix  $W = (w_{ij})$  by

$$w_{ij} = \begin{cases} \alpha & \text{if } \{i, j\} \in E \\ -\alpha & \text{if } \{i, j\} \notin E, \end{cases}$$

where  $0 < \alpha < 1$  is a parameter.

We consider the polynomial  $h(z)$  defined by (4.2.1) and let  $f(z) = \ln h(z)$ . Our goal is to approximate  $f(1) = \ln h(1)$  and hence

$$\begin{aligned} h(1) &= \sum_{\substack{S \subset \{1, \dots, n\} \\ |S|=m}} (1 + \alpha)^{\binom{m}{2} \sigma(S)} (1 - \alpha)^{\binom{m}{2} (1 - \sigma(S))} \\ &= (1 - \alpha)^{\binom{m}{2}} \text{den}_m(G; \gamma), \quad \text{where} \quad \gamma = \frac{m-1}{2} \ln \frac{1+\alpha}{1-\alpha}. \end{aligned}$$

We approximate  $f(1)$  by the degree  $r$  Taylor polynomial of  $f(z)$  computed at  $z = 0$ . The results of [Ba15] suggest that for  $\alpha = O(1/m)$ , we should get a reasonable approximation if we use  $r \sim \ln m$ . The results of our numerical experiments suggest that we get reasonable approximations if we use  $\alpha = \Omega(1)$  and  $r = 2$  or  $r = 3$ . In short, on the examples we tested, the quality of approximation was more consistent with the quality of the Taylor polynomial approximation of  $\ln(1 \pm \alpha)$ .

More precisely, we ran the algorithm typically with parameters  $n = 50, 100$  and  $m = 10$ , although occasionally we chose  $n$  as large as  $n = 300$ . For the parameters  $n = 50$  and  $m = 10$  we were able to compare our approximation with the exact value. Typically, choosing  $\alpha = 0.5$  or lower produced an approximation of  $f(1)$  within 1% accuracy. For  $\alpha = 0.7$ , the accuracy went down to 10% – 20% and for  $\alpha > 0.7$  the approximation was not accurate. For higher values of  $n$ , where the exact value of  $f(1)$  was unavailable, we compared the approximations obtained for  $r = 2$  and  $r = 3$ . If the approximations were close to each other, we considered it as an indication that they are also close to the true value of  $f(1)$ . Again, we observed that up to  $\alpha = 0.5$ , the approximations agreed, but were beginning to essentially differ at  $\alpha = 0.7$  and higher. For the graphs, we used the Erdős -Rényi models  $G(n, 0.5)$ ,  $G(n, 0.4)$ , those graphs with planted cliques of size  $m$ , and occasionally manually constructed “random-looking” graphs.

We provide below the explicit formulas for the approximations up to degree 3, in case the reader will be interested to do some numerical experiments. We interpret  $w_{ij}$  as weights on the edges of a complete graph with  $n$  vertices. Borrowing an idea from [PR17], we express the derivatives  $f^{(k)}(0)$  in terms of various sums associated with *connected* subgraphs, since it improves the computational complexity of the algorithm. We remark, however, that it looks unlikely that the methods of [PR17] can be pushed to improve the complexity of our algorithm in the general situation from quasi-polynomial to genuinely polynomial, since we work with graphs of unbounded degrees.

It is convenient to introduce the following sums:

$$A_1 = \sum_{\{i,j\}} w_{ij},$$

where the sum is taken over all unordered pairs  $\{i, j\}$  of distinct indices;

$$B_1 = \sum_{\{i,j\}} w_{ij}^2, \quad B_2 = \sum_{j, \{i,k\}} w_{ij} w_{jk},$$

where in the formula for  $B_1$  the sum is taken over all unordered pairs  $\{i, j\}$  of distinct indices and in  $B_2$  the sum is taken over all pairs consisting of an index  $j$  and an unordered pair  $\{i, k\}$ , so that all three indices are distinct; and

$$C_1 = \sum_{\{i,j\}} w_{ij}^3, \quad C_2 = \sum_{(i,j,k)} w_{ij}^2 w_{jk}, \quad C_3 = \sum_{\{i,j,k\}} w_{ij} w_{jk} w_{ki},$$

$$C_4 = \sum_{(i,j,k,l)} w_{ij} w_{jk} w_{kl}, \quad C_5 = \sum_{\{j,k,l\},i} w_{il} w_{ij} w_{ik},$$

where in  $C_1$  the sum is taken over all unordered pairs  $\{i, j\}$  of distinct indices, in  $C_2$  the sum is taken over all ordered triples  $(i, j, k)$  of distinct indices, in  $C_3$  the sum is taken over all unordered triples of distinct integers, in  $C_4$ , the sum is taken over all ordered 4-tuples  $(i, j, k, l)$  of distinct indices, and in  $C_5$  the sum is taken over all pairs consisting of an index  $i$  and an unordered triple  $\{j, k, l\}$  so that all four indices  $\{i, j, k, l\}$  are distinct.

**(5.1) First-order approximation.** Clearly,  $h(0) = 1$ . From (4.4.2), we have

$$h'(0) = \binom{n}{m}^{-1} \binom{n-2}{m-2} \sum_{\{i,j\} \subset \{1,\dots,n\}} w_{ij} = \frac{m(m-1)}{n(n-1)} A_1.$$

Since  $f(0) = \ln h(0) = 0$  and  $f'(0) = h'(0)/h(0) = h'(0)$ , we obtain the first order approximation

$$f(1) \approx h'(0),$$

where  $h'(0)$  is defined as above. The complexity of computing the first order approximation is  $O(n^2)$ .

**(5.2) Second-order approximation.** From (4.4.2), we have

$$h''(0) = \binom{n}{m}^{-1} \sum_{I=\{\{i_1,j_1\},\{i_2,j_2\}\}} \binom{n-\nu(I)}{m-\nu(I)} w_{i_1 j_1} w_{i_2 j_2}.$$

Here  $\nu(I) = 4$  if the pairs  $\{i_1, j_1\}$  and  $\{i_2, j_2\}$  are pairwise disjoint and  $\nu(I) = 3$  if they share exactly one index. Hence we can write

$$h''(0) = \binom{n}{m}^{-1} \left( 2 \binom{n-3}{m-3} B_2 + \binom{n-4}{m-4} (A_1^2 - 2B_2 - B_1) \right)$$

$$= 2 \frac{m(m-1)(m-2)}{n(n-1)(n-2)} B_2 + \frac{m(m-1)(m-2)(m-3)}{n(n-1)(n-2)(n-3)} (A_1^2 - 2B_2 - B_1).$$

Since

$$f''(0) = h''(0) - (h'(0))^2,$$

we obtain the second order approximation:

$$f(1) \approx f'(0) + \frac{1}{2} f''(0) = h'(0) - \frac{1}{2} (h'(0))^2 + \frac{1}{2} h''(0),$$

where  $h'(0)$  and  $h''(0)$  are defined as above. The complexity of computing the second order approximation is  $O(n^3)$ .

**(5.3) Third-order approximation.** From (4.4.2), one can deduce that

$$\begin{aligned}
h'''(0) = & 6 \frac{m(m-1)(m-2)}{n(n-1)(n-2)} C_3 + \frac{m(m-1)(m-2)(m-3)}{n(n-1)(n-2)(n-3)} (6C_5 + 3C_4) \\
& + 6 \frac{m(m-1)(m-2)(m-3)(m-4)}{n(n-1)(n-2)(n-3)(n-4)} (A_1 B_2 - 3C_5 - 3C_3 - C_4 - C_2) \\
& + \frac{m(m-1)(m-2)(m-3)(m-4)(m-5)}{n(n-1)(n-2)(n-3)(n-4)(n-5)} \left( A_1^3 + 12C_3 - 6A_1 B_2 \right. \\
& \quad \left. + 12C_5 + 3C_4 + 6C_2 - 3A_1 B_1 + 2C_1 \right).
\end{aligned}$$

Since we have

$$f'''(0) = h'''(0) - 2f''(0)h'(0) - f'(0)h''(0) = 2(h'(0))^3 - 3h'(0)h''(0) + h'''(0),$$

we obtain the third order approximation

$$\begin{aligned}
f(1) & \approx f'(0) + \frac{1}{2}f''(0) + \frac{1}{6}f'''(0) \\
& = h'(0) - \frac{1}{2}(h'(0))^2 + \frac{1}{2}h''(0) + \frac{1}{3}(h'(0))^3 - \frac{1}{2}h'(0)h''(0) + \frac{1}{6}h'''(0).
\end{aligned}$$

The complexity of computing the third order approximation is  $O(n^4)$ .

## 6. PROOF OF THEOREM 1.5 AND CONCLUDING REMARKS

We got the idea of the proof from [EM18], where a similar question about complex zeros of the permanents of matrices with independent random entries was treated.

Applying Jensen's formula, see for example, Section 5.3 of [Ah78], we obtain

$$(6.1) \quad \ln |h_W(0)| = \sum_{s=1}^N \ln \frac{|a_{s,W}|}{r} + \frac{1}{2\pi} \int_0^{2\pi} \ln |h_W(re^{i\theta})| d\theta,$$

where  $a_{s,W}$ ,  $s = 1, \dots, N$  are the roots of the polynomial  $h_W(z)$  in the disc  $|z| < r$  and we assume that  $h_W(z)$  has no zeros on the circle  $|z| = r$  (since there are only finitely many values of  $r$  with roots on the circle  $|z| = r$ , this assumption is not restrictive). We have

$$\ln |h_W(0)| = 0$$

and furthermore, applying Jensen's inequality, we bound:

$$\begin{aligned}
(6.2) \quad \frac{1}{2\pi} \int_0^{2\pi} \ln |h_W(re^{i\theta})| d\theta & = \frac{1}{2} \cdot \frac{1}{2\pi} \int_0^{2\pi} \ln |h_W(re^{i\theta})|^2 d\theta \\
& \leq \frac{1}{2} \ln \left( \frac{1}{2\pi} \int_0^{2\pi} |h_W(re^{i\theta})|^2 d\theta \right).
\end{aligned}$$

For a fixed  $\theta \in [0, 2\pi]$ , we compute the expectation

$$\begin{aligned} \mathbf{E} |h_W(re^{i\theta})|^2 &= \binom{n}{m}^{-2} \sum_{\substack{S_1, S_2 \subset \{1, \dots, n\} \\ |S_1| = |S_2| = m}} \mathbf{E} \left( \prod_{\{j, k\} \subset S_1} (1 + re^{i\theta} w_{jk}) \right. \\ &\quad \left. \times \prod_{\{j, k\} \subset S_2} (1 + re^{-i\theta} w_{jk}) \right) \\ &= \binom{n}{m}^{-2} \sum_{\substack{S_1, S_2 \subset \{1, \dots, n\} \\ |S_1| = |S_2| = m}} (1 + r^2)^{\binom{|S_1 \cap S_2|}{2}}. \end{aligned}$$

A subset  $S \subset \{1, \dots, n\}$  of cardinality  $l = |S| \leq m$  can be represented as the intersection  $S = S_1 \cap S_2$  of  $m$ -subsets  $S_1, S_2$  in  $\binom{n-l}{m-l} \binom{n-m}{m-l}$  ways. Hence

$$(6.3) \quad \mathbf{E} |h_W(re^{i\theta})|^2 = \binom{n}{m}^{-2} \sum_{l=0}^m \binom{n}{l} \binom{n-l}{m-l} \binom{n-m}{m-l} (1 + r^2)^{\binom{l}{2}}.$$

To bound (6.3), we consider the ratio of the  $(l+1)$ -st term to the  $l$ -th term:

$$\begin{aligned} \frac{n-l}{l+1} \cdot \frac{m-l}{n-l} \cdot \frac{m-l}{n-2m+l+1} \cdot (1+r^2)^l &= \frac{(m-l)^2 (1+r^2)^l}{(l+1)(n-2m+l+1)} \\ &\leq \frac{m^2(1+r^2)^m}{n-2m+1}. \end{aligned}$$

In particular, if

$$(6.4) \quad n \geq 2m^2(1+r^2)^m + 2m,$$

the ratio does not exceed  $1/2$  and hence we can bound the sum (6.3) by

$$\mathbf{E} |h_W(re^{i\theta})|^2 \leq 2 \binom{n}{m}^{-2} \binom{n}{m} \binom{n-m}{m} \leq 2.$$

Integrating over  $\theta$ , we conclude that if (6.4) holds then

$$\mathbf{E} \left( \frac{1}{2\pi} \int_0^{2\pi} |h_W(re^{i\theta})| d\theta \right) \leq 2.$$

By the Markov inequality, for any  $\tau \geq 1$ , we get

$$\mathbf{P} \left( \frac{1}{2\pi} \int_0^{2\pi} |h_W(re^{i\theta})| d\theta \geq 2\tau \right) \leq \frac{1}{\tau}.$$

Consequently, from (6.1) and (6.2), we have

$$\mathbf{P} \left( \sum_{s=1}^N \ln \frac{|a_{s,W}|}{r} \leq -\frac{1}{2} \ln 2\tau \right) \leq \frac{1}{\tau}.$$

and the proof follows.  $\square$

An anonymous referee asked what happens if  $G$  is a random graph  $G(n, 0.5)$  with a planted  $m$ -clique. The most interesting asymptotic regime is when  $m^2 \ll n \leq m^{O(1)}$  and  $m$  grows, see [A+98] for results and references. Here we are interested in a polynomial time algorithm which, with high probability, tells  $G$  from  $G(n, 0.5)$ . A quasi-polynomial time algorithm is readily available (by an exhaustive search for a clique of size at least  $3 \log_2 n$ , say). Our proof of Theorem 1.5 does not seem to extend to random graphs with a planted clique. We note, however, that if the radius of zero-free region is roughly the same  $r = \Omega(1/\sqrt{m})$  as in Theorem 1.5 or even weaker,  $r = \Omega(m^{-1+\epsilon})$  for some  $\epsilon > 0$ , we do obtain a desired polynomial time algorithm. Indeed, in the latter case, we can choose  $\gamma = m^{\epsilon'}$  with some  $0 < \epsilon' < \epsilon$ . If  $G$  is a graph with a planted  $m$ -clique, we have

$$\text{den}_m(G; \gamma) \geq \exp \left\{ m^{1+\epsilon'} - O(m \ln m) \right\},$$

cf. (1.1.2). If  $G$  is a random graph  $G(n, 0.5)$ , our proof Theorem 1.5 implies that

$$\text{den}_m(G; \gamma) \leq \exp \left\{ \frac{m^{1+\epsilon'}}{2} + O(1) \right\}$$

with high probability, cf. (1.5.1). Note that by choosing  $\epsilon' < \epsilon$ , we choose  $\gamma$  sufficiently “deep” inside the purported zero-free region, and hence we can get a genuinely polynomial, as opposed to a quasi-polynomial, algorithm by computing a constant, as opposed to logarithmic, number of terms in the Taylor polynomial approximation, cf. Lemma 4.2.

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